ECE 20875 Python for Data Science

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(Adapted from material developed by Profs. Milind Kulkarni, Stanley Chan, Chris Brinton, David Inouye)

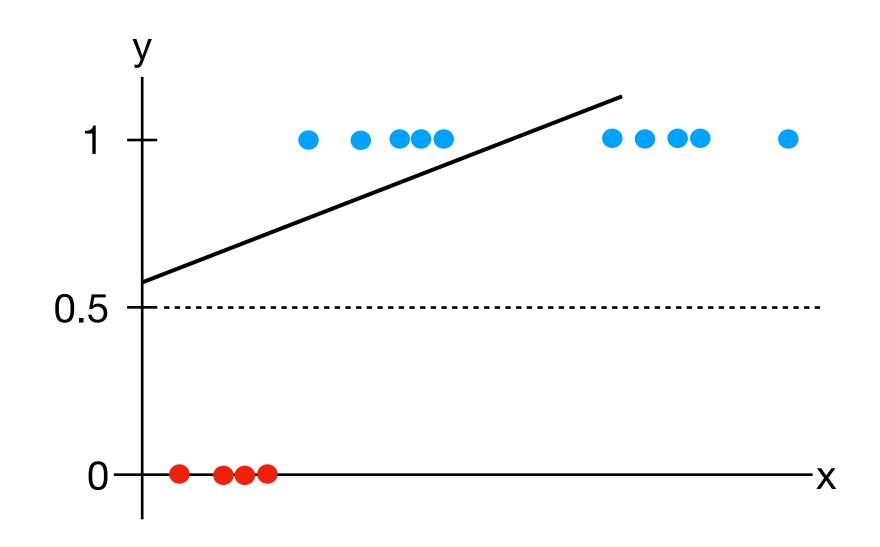
classification: logistic regression

regression with two classes

• With linear regression, we model the relationship between features and target with a linear equation:

$$\hat{y}_i = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_m x_m$$

- Now, suppose we have two classes, i.e., $y \in \{0, 1\}$. We could use linear regression, but ...
 - it will treat the classes as numbers, interpolating between the points
 - it cannot be interpreted as a probability
 - how would we generalize to multiple classes?



- Need a decision threshold, i.e., y = 0.5
- In this case, we would never predict the class y = 0, regardless of what x is!

logistic regression model

 Instead of fitting a hyperplane (a line generalized to more than one dimension), use the logistic function

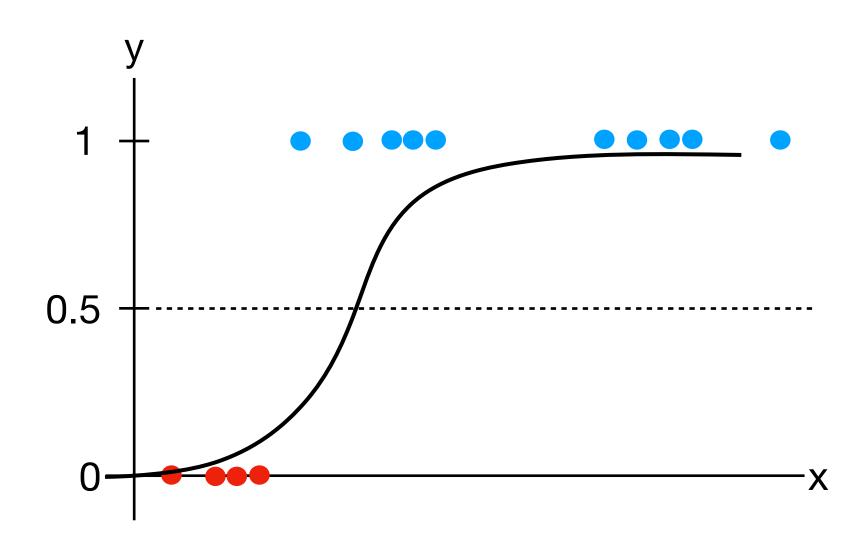
$$g(v) = \frac{1}{1 + e^{-v}}$$

to translate the output of linear regression to between 0 (as $v \to -\infty$) and 1 (as $v \to \infty$)

- Also note that $1 g(v) = \frac{e^{-v}}{1 + e^{-v}}$ (useful for derivations)
- This converts the outputs to probabilities:

$$f_{\beta}(x) = g(\beta_0 + \beta^T x) = P(y = 1 \mid x)$$

$$= \frac{1}{1 + \exp(-(\beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_m x_m))}$$



- Now the decision rule
 - $\hat{y}(x) \ge 0.5 \to \hat{y} = 1$
 - $\hat{y}(x) < 0.5 \rightarrow \hat{y} = 0$

has a probabilistic interpretation

interpreting coefficients

- In linear regression, the effect of a coefficient is clear: $\beta_j x_j$ means for every unit change in x_j , the model changes by β_j
- For logistic regression, we need to find a different interpretation, since the weights no longer have a linear effect
- Consider the **odds**, i.e., the ratio P(y = 1 | x)/P(y = 0 | x):

$$\frac{P(y=1|x)}{P(y=0|x)} = \frac{1}{1 + \exp(-(\beta_0 + \beta_1 x_1 + \dots + \beta_m x_m))} \cdot \frac{1 + \exp(-(\beta_0 + \beta_1 x_1 + \dots + \beta_m x_m))}{\exp(-(\beta_0 + \beta_1 x_1 + \dots + \beta_m x_m))}$$

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$$= \frac{1}{\exp(-(\beta_0 + \beta_1 x_1 + \dots + \beta_m x_m))} = \exp(\beta_0 + \beta_1 x_1 + \dots + \beta_m x_m)$$

interpreting coefficients

• Then we consider the ratio of the odds when x_i is increased by 1:

$$\frac{\text{odds}_{x_j+1}}{\text{odds}_{x_j}} = \frac{\exp(\cdots + \beta_j(x_j+1) + \cdots)}{\exp(\cdots + \beta_j x_j + \cdots)} = e^{\beta_j}$$

- Thus, a unit change in x_{ij} corresponds to a factor e^{β_j} change in the odds
 - $e^{\beta_j} > 1$: x_j increases the odds
 - $e^{\beta_j} < 1$: x_j decreases the odds

Consider

$$\hat{y} = \frac{1}{1 + \exp(-(3 + 2x_1 + 0.5x_2 - 3x_3))}$$

- For this model
 - x_1 and x_2 increase the odds
 - x_3 decreases the odds
 - x_3 has the largest factor impact on the odds (assuming the features are normalized!)

training logistic regression

- With linear regression, we can derive a closed-form solution for the parameters in terms of the least-squares equations
- For logistic regression, let's consider the likelihood of the model over data samples i = 1,...,n:

$$L(\beta) = \prod_{i=1}^n p(y_i \mid x_i, \beta) = \prod_{i=1}^n (f_\beta(x_i))^{y_i} \cdot (1 - f_\beta(x_i))^{1 - y_i} \qquad \text{when } y_i = 1 \text{, we want to maximize } f_\beta(x_i) \text{, and when } y_i = 0 \text{, we want to maximize } 1 - f_\beta(x_i)$$

And then the log likelihood, which is easier to optimize (like we did with GMMs):

$$l(\beta) = \sum_{i=1}^{n} \log \left[(f_{\beta}(x_i))^{y_i} \cdot (1 - f_{\beta}(x_i))^{1 - y_i} \right] = \sum_{i=1}^{n} \left[y_i \log f_{\beta}(x_i) + (1 - y_i) \log(1 - f_{\beta}(x_i)) \right]$$

• There is no (known) closed form solution to maximize $l(\beta)$, given the $\log f_{\beta}(x_i)$ terms

gradient descent (ascent)

- We want to find β to maximize $l(\beta)$
- Consider the **gradient descent (ascent)** algorithm, an iterative procedure for finding a **local minimum (maximum)** of a function by moving away from (towards) the gradient:

$$\beta_j^{t+1} = \beta_j^t - \alpha^t \frac{\partial}{\partial \beta_j} l(\beta^t), \qquad \beta_j^{t+1} = \beta_j^t + \alpha^t \frac{\partial}{\partial \beta_j} l(\beta^t)$$

- Here, α^t is the **step size** of the algorithm at time t
- Since $l(\beta)$ is a **concave** function, we can *guarantee* that gradient ascent will eventually converge to the **global** maximum, so long as certain conditions on α^t are met

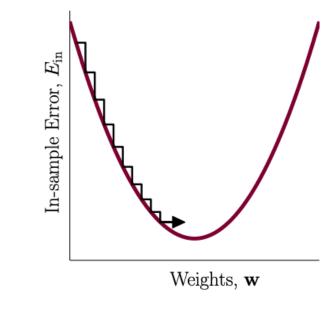
for non-convex functions, no guarantee of convergence to optimum

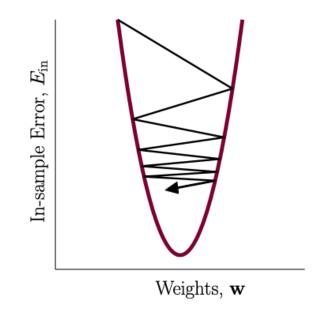


in general, the step size must be tuned correctly

 η too small

 η too large





example

Suppose we have a single parameter b for some model we are trying to train. For this model, we find a log-likelihood function of

$$l(b) = -\left(\frac{b-m}{s}\right)^2$$

where m and s are constants. Derive the iterative procedure for determining the model parameters as a function of the step size α^t . Run the procedure for different values of α^t until t=10 and compare the results.

solution

We always want to maximize the loglikelihood, so we use gradient ascent. Letting b^t be the value of b at iteration t, our update procedure will be:

$$b^{t+1} = b^t + \alpha^t \frac{d}{db} l(b^t)$$

Evaluating the derivative, this becomes:

$$b^{t+1} = b^t - 2\alpha^t \left(\frac{b^t - m}{s^2}\right)$$

Suppose $\alpha=0.1, m=5, s=0.7$. If we start at $b^0=1.1$ (arbitrary), we get

$$b^{1} = 1.1 - 2 \cdot 0.1 \cdot \left(\frac{1.1 - 5}{0.7^{2}}\right) = 2.692$$

$$b^2 = 2.692 - 2 \cdot 0.1 \cdot \left(\frac{2.692 - 5}{0.7^2}\right) = 3.634$$

. . .

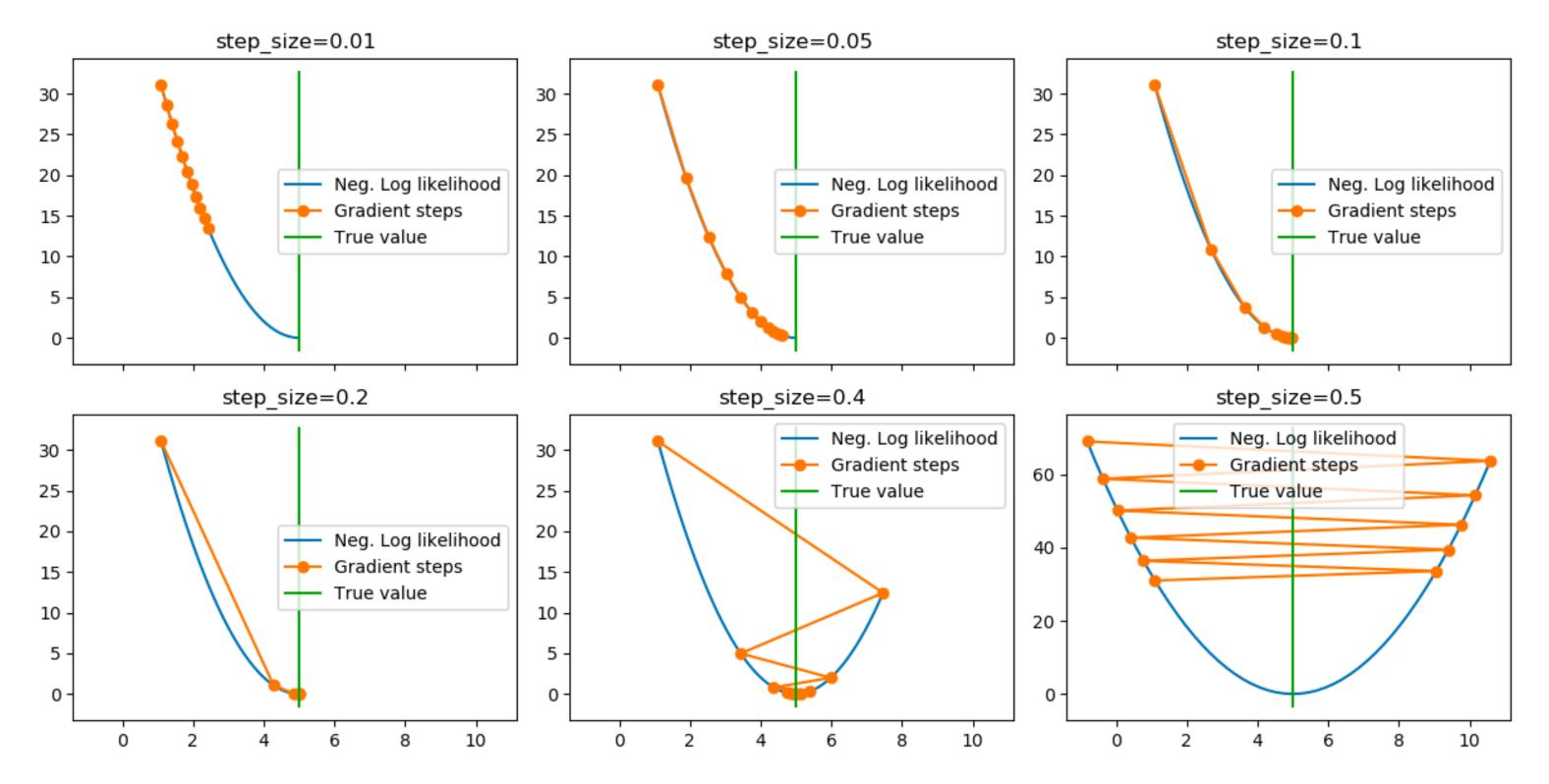
$$b^{10} = 4.965 - 2 \cdot 0.1 \cdot \left(\frac{4.965 - 5}{0.7^2}\right) = 4.979$$

solution

Below, we plot the evolution of b^t over t (see the Jupyter notebook), starting with $b^0 = 1.1$ for $\alpha^t = 0.01, 0.05, 0.1, 0.2, 0.4, 0.5$. Again, we set m = 5 and s = 0.7.

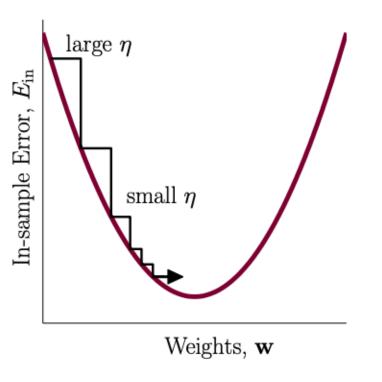
Here, the y-axis is actually -l(b), to make the values positive. Maximizing the log-likelihood is equivalent to minimizing the negative log-likelihood.

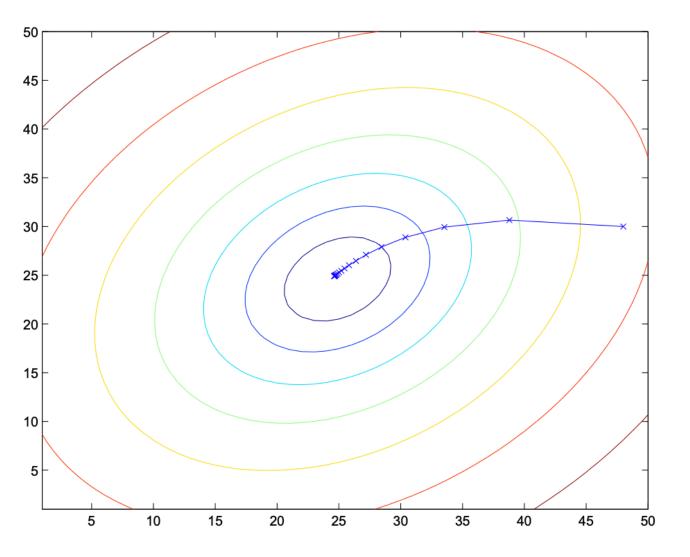
Tuning α^t is a very important question!



• Back to logistic regression. Evaluating the partial derivative,

$$\frac{\partial}{\partial \beta_j} l(\beta) = \sum_{i=1}^n \left(\frac{y_i}{f_{\beta}(x_i)} - \frac{1 - y_i}{1 - f_{\beta}(x_i)} \right) \frac{\partial}{\partial \beta_j} f_{\beta}(x_i)$$

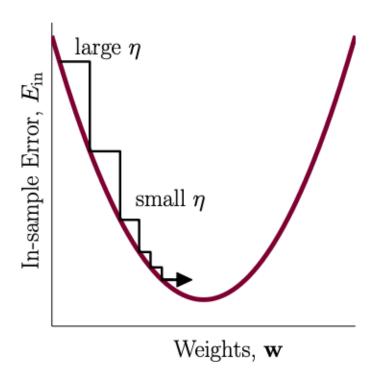


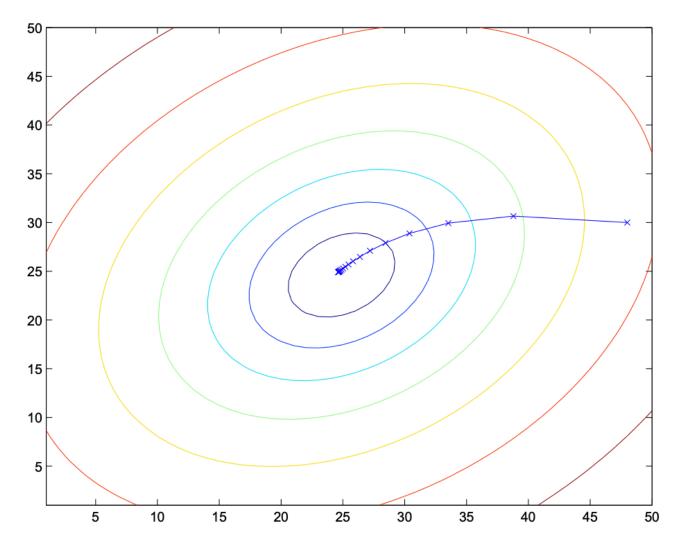


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$$= \sum_{i=1}^{n} \left(\frac{y_i}{f_{\beta}(x_i)} - \frac{1 - y_i}{1 - f_{\beta}(x_i)} \right) f_{\beta}(x_i) (1 - f_{\beta}(x_i)) \frac{\partial}{\partial \beta_j} (\dots + \beta_j x_{ij} + \dots)$$



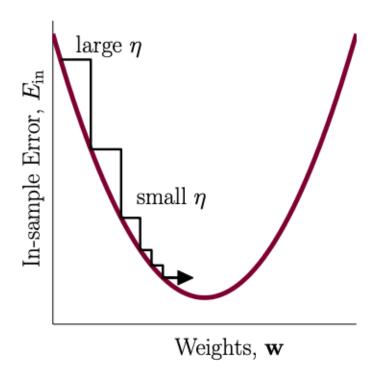


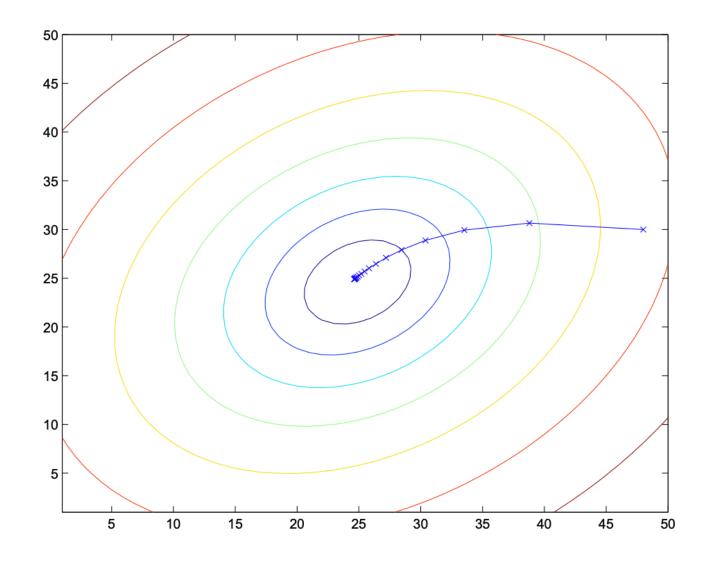
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$$= \sum_{i=1}^{n} \left(y_i (1 - f_{\beta}(x_i)) - (1 - y_i) f_{\beta}(x_i) \right) x_{ij} = \sum_{i=1}^{m} (y_i - f_{\beta}(x_i)) x_{ij}$$





• Back to logistic regression. Evaluating the partial derivative,

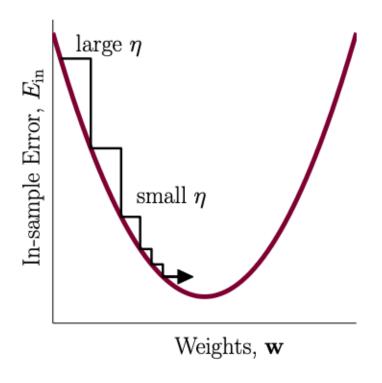
$$\frac{\partial}{\partial \beta_j} l(\beta) = \sum_{i=1}^n \left(\frac{y_i}{f_{\beta}(x_i)} - \frac{1 - y_i}{1 - f_{\beta}(x_i)} \right) \frac{\partial}{\partial \beta_j} f_{\beta}(x_i)$$

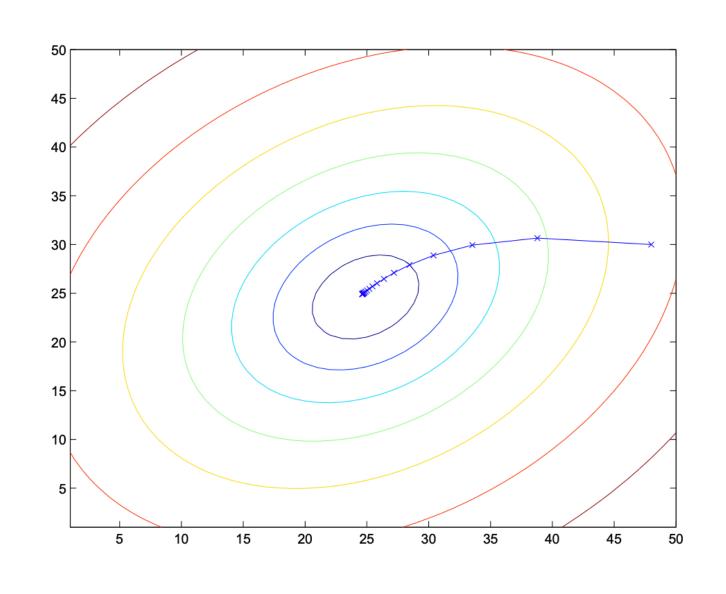
$$= \sum_{i=1}^{n} \left(\frac{y_i}{f_{\beta}(x_i)} - \frac{1 - y_i}{1 - f_{\beta}(x_i)} \right) f_{\beta}(x_i) (1 - f_{\beta}(x_i)) \frac{\partial}{\partial \beta_j} (\dots + \beta_j x_{ij} + \dots)$$

$$= \sum_{i=1}^{n} \left(y_i (1 - f_{\beta}(x_i)) - (1 - y_i) f_{\beta}(x_i) \right) x_{ij} = \sum_{i=1}^{m} (y_i - f_{\beta}(x_i)) x_{ij}$$

• Thus, we get the following gradient ascent rule for logistic regression:

$$\beta_j^{t+1} = \beta_j^t + \alpha^t \left[\sum_{i=1}^n (y_i - f_{\beta}(x_i)) x_{ij} \right]$$





in python

_pred)

- from sklearn.linear_model import LogisticRegression
 - https://scikit-learn.org/stable/ modules/generated/ sklearn.linear_model.LogisticRegres sion.html
- Most methods (fit, predict, ...) are the same as linear regression
- One difference: Regularization parameter C
 - Higher *C*: Less regularization
 - Lower *C*: *More* regularization

```
from sklearn.linear model
import LogisticRegression
from sklearn import metrics
logreg = LogisticRegression()
logreg.fit(X_train,y_train)
y_pred = logreg.predict(X_test)
metrics.accuracy score(y test,y
```