

ECE 20875

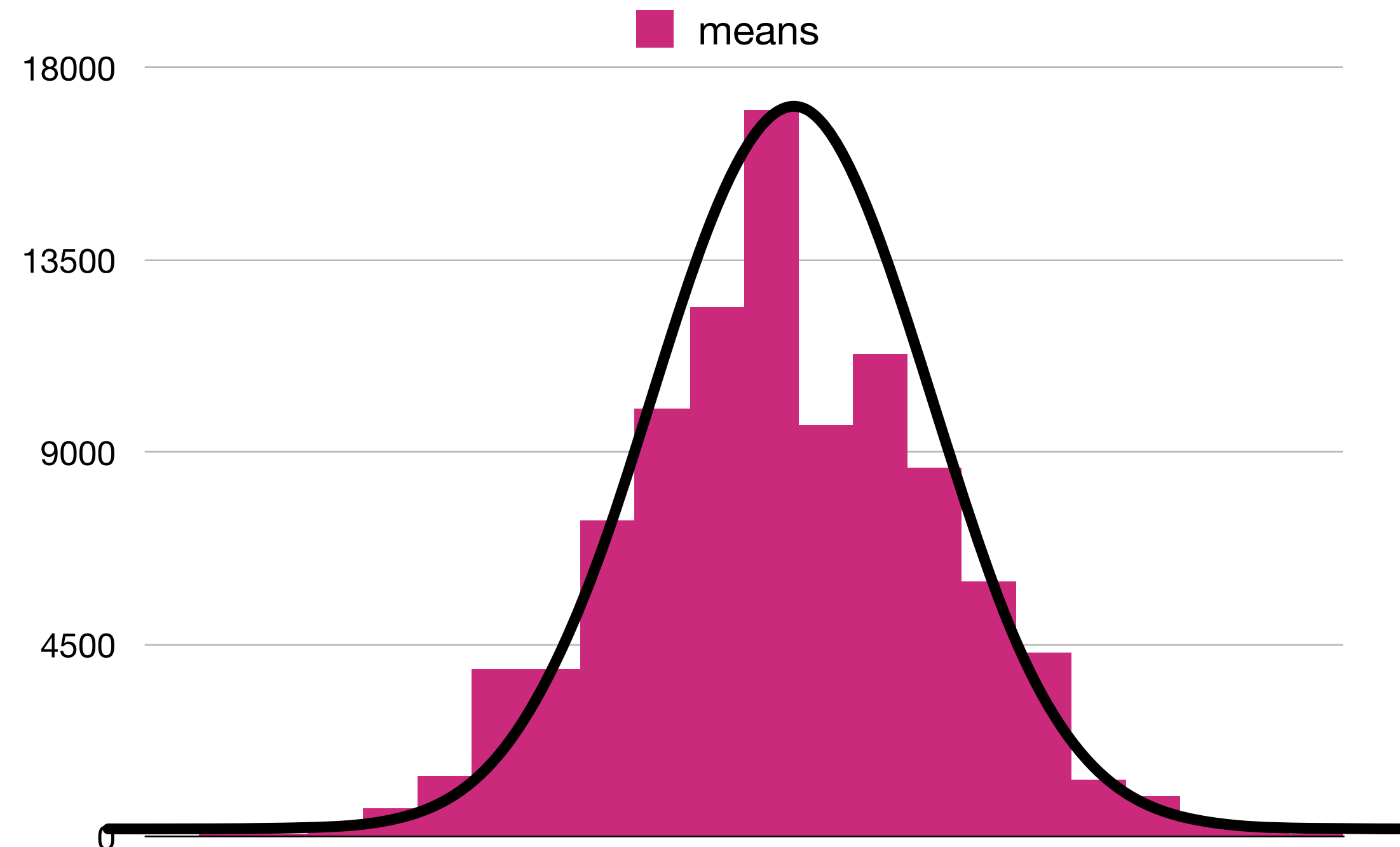
Python for Data Science

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**(Adapted from material developed by Profs. Milind Kulkarni,
Stanley Chan, Chris Brinton, David Inouye, Qiang Qiu)**

**confidence intervals and
hypothesis testing**

sampling distribution

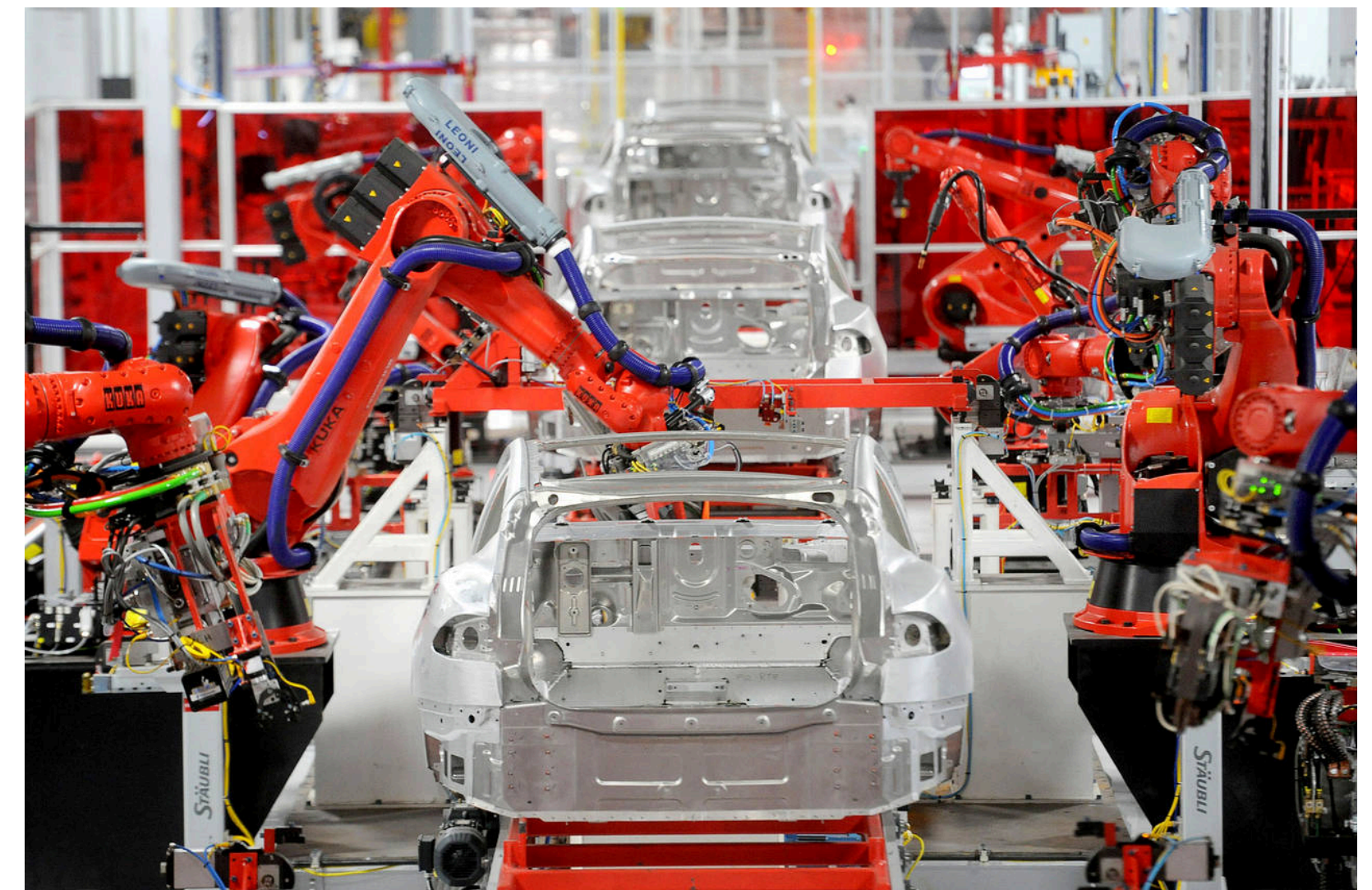


Each data point is the \bar{x} of one experiment

- Recall that by the central limit theorem, sample means approach a normal distribution
- Can we use this to draw conclusions about our data?

asking questions about data

- Suppose a factory claims to produce widgets with an average weight of 100g and a standard deviation of 22g
- We receive a new shipment of widgets which seem off, and we want to see whether the factory has shifted
- Form two hypotheses:
 - **Null hypothesis (H_0):** The factory is producing according to specification, i.e., $\mu = 100g$.
 - **Alternative hypothesis (H_1):** The factory is not producing according to specification, i.e., $\mu \neq 100g$.
- Suppose we weigh 100 of the new widgets (i.e., sample $n = 100$ widgets) and find their average weight is $\bar{x} = 95g$
 - What can we conclude?

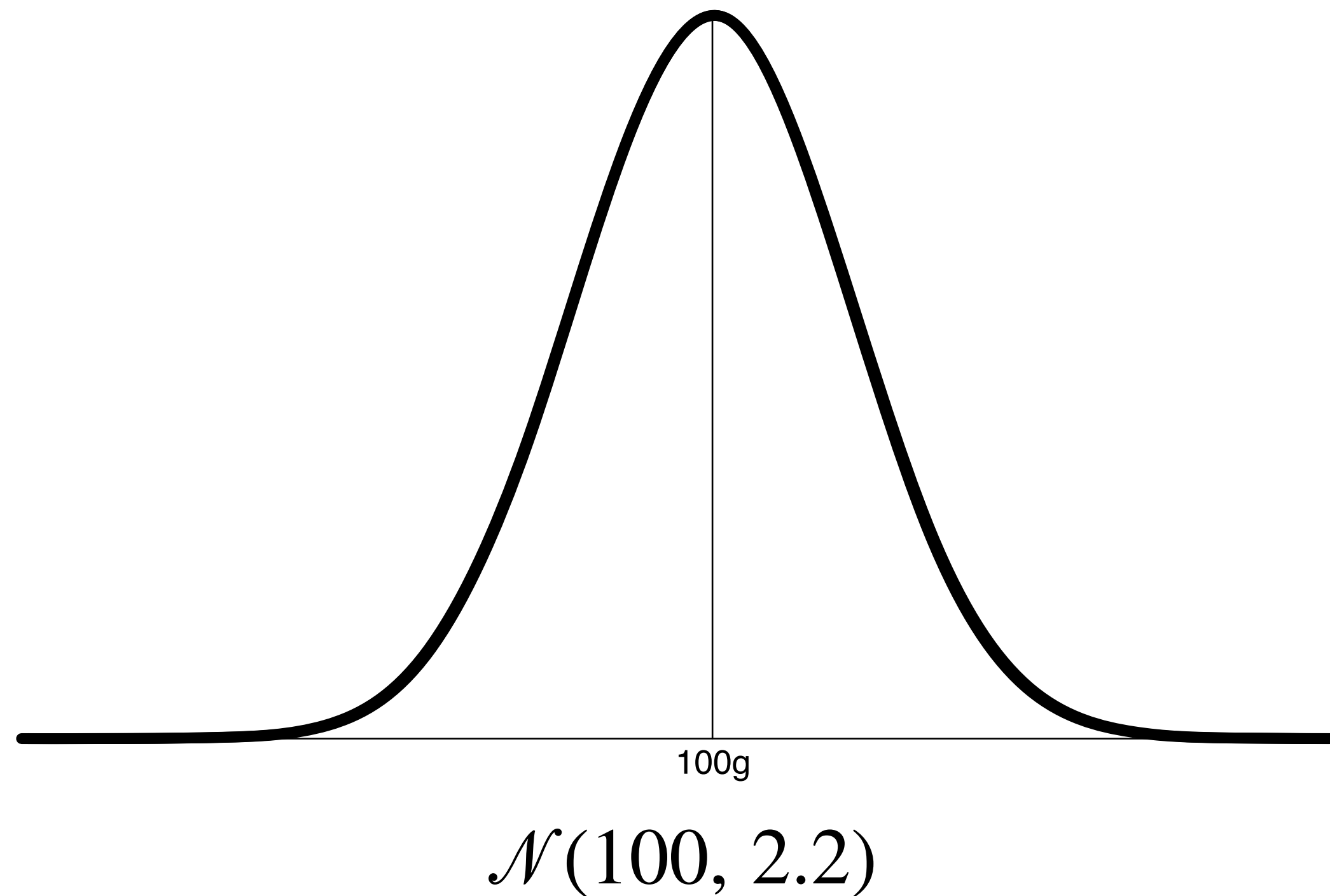


asking questions about data

- **Are the widgets in spec?**
- Not as simple as it seems!
- We have picked one sample of widgets, but it could just be a bad sample!
- Can we use our sampling distribution to help?



hypothesis testing

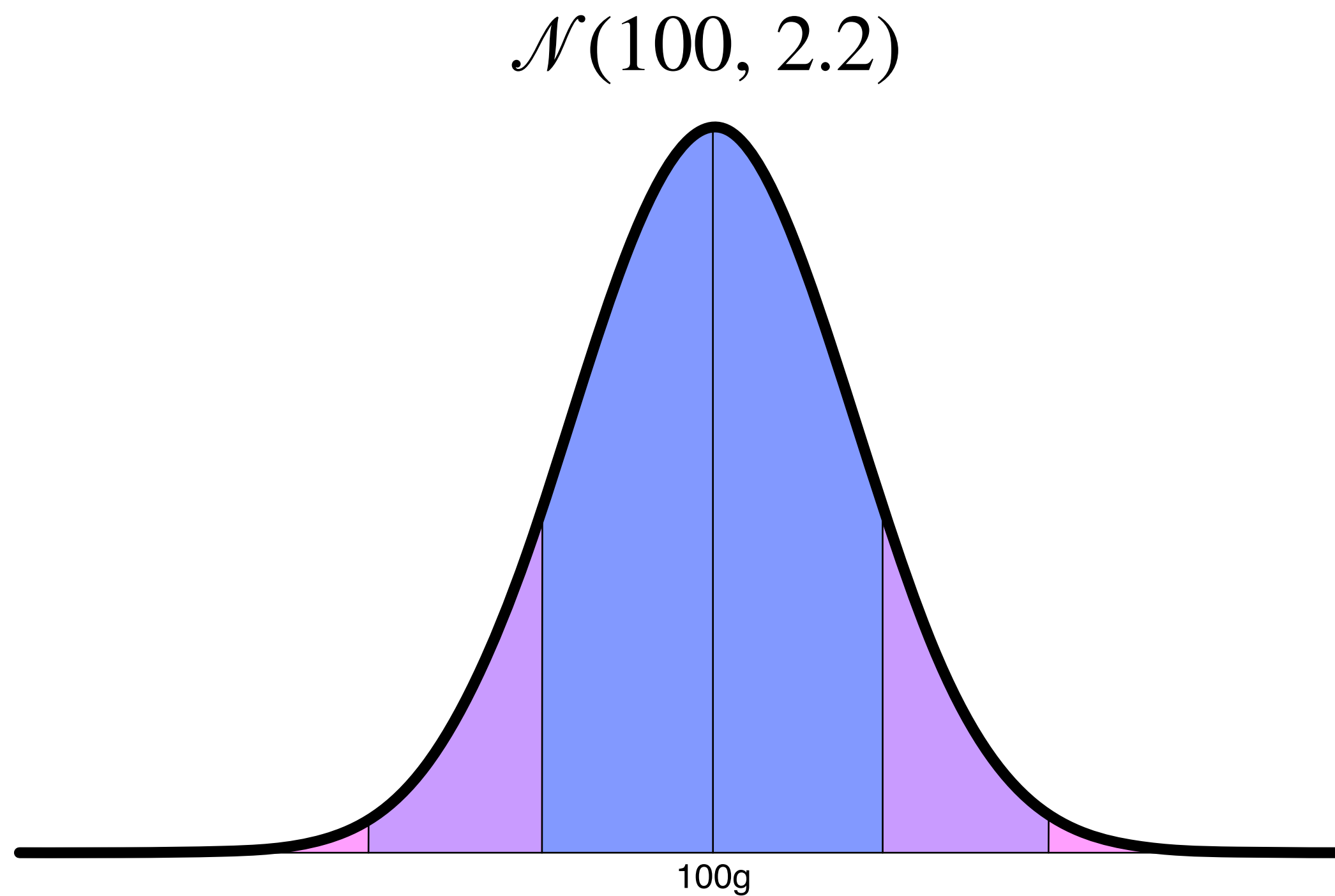


- Suppose the null hypothesis is true (new widgets are from the same distribution as the original widgets)
- Then the sampling distribution should have its mean at $\mu = 100\text{g}$
- And the sampling distribution should have a standard deviation of:

$$SE \triangleq \sigma_{\bar{X}} = \frac{\sigma}{\sqrt{n}} \approx \frac{22}{10} = 2.2\text{g}$$

- This is called the **standard error** (SE)
- Remember, σ is from the population, which we sometimes have to estimate with s (from the sample)

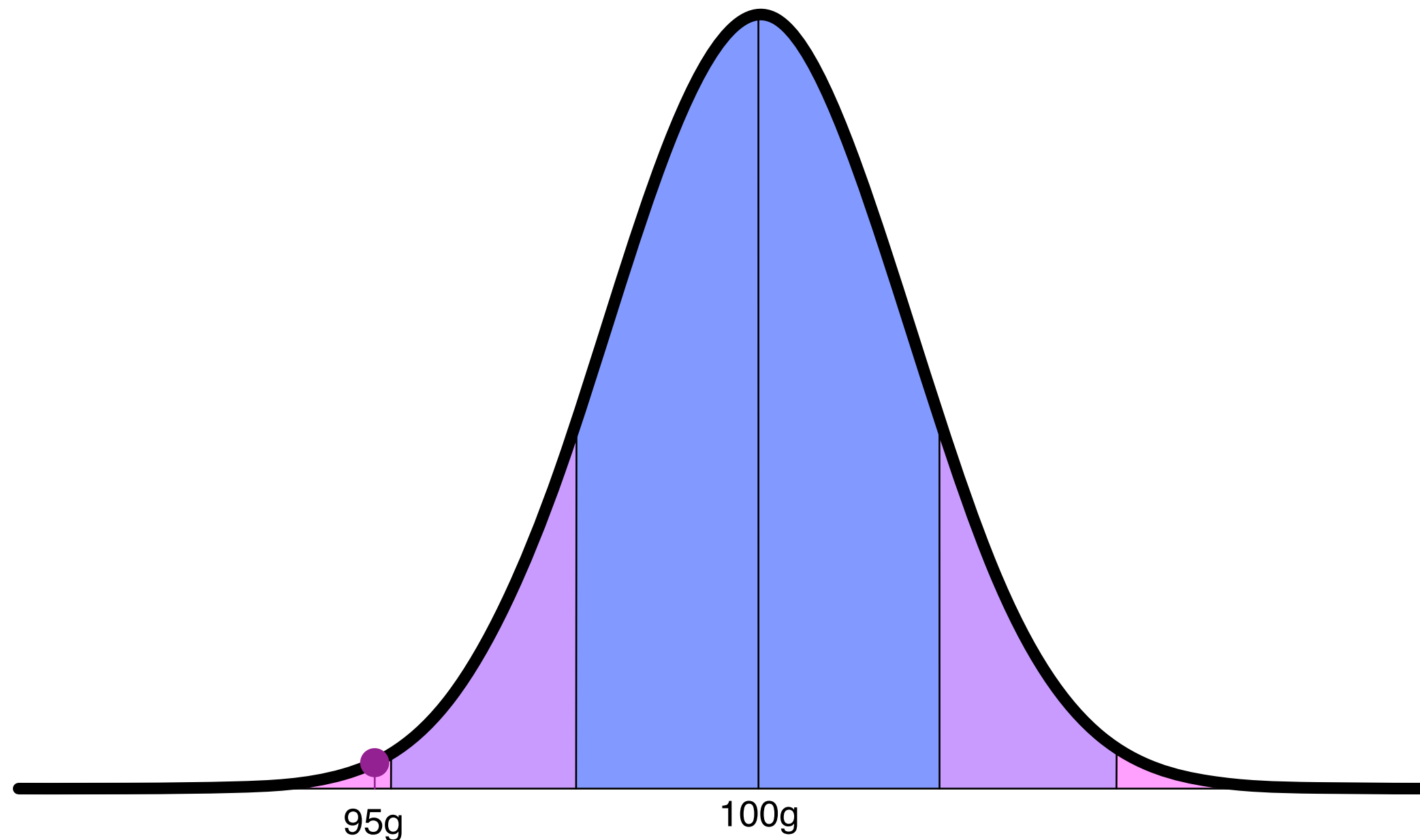
hypothesis testing



- Remember properties of normal distribution:
 - ~68% of points within one σ of μ
 - ~95% of points within two σ of μ
 - ~99.7% of points within three σ of μ

hypothesis testing

- So what about our sample \bar{x} of 95g?



- Very unlikely for it to have come from this distribution!

- Remember properties of normal distribution:
 - ~68% of points within one σ of μ
 - ~95% of points within two σ of μ
 - ~99.7% of points within three σ of μ
- 95g is between 2 and 3 $\sigma_{\bar{X}}$ of μ

z-test

- The statistical **z-test**

- Reasoning about μ

- Applicable when we know σ or if n is large enough (if we don't know σ and n is large enough, we can estimate with s)

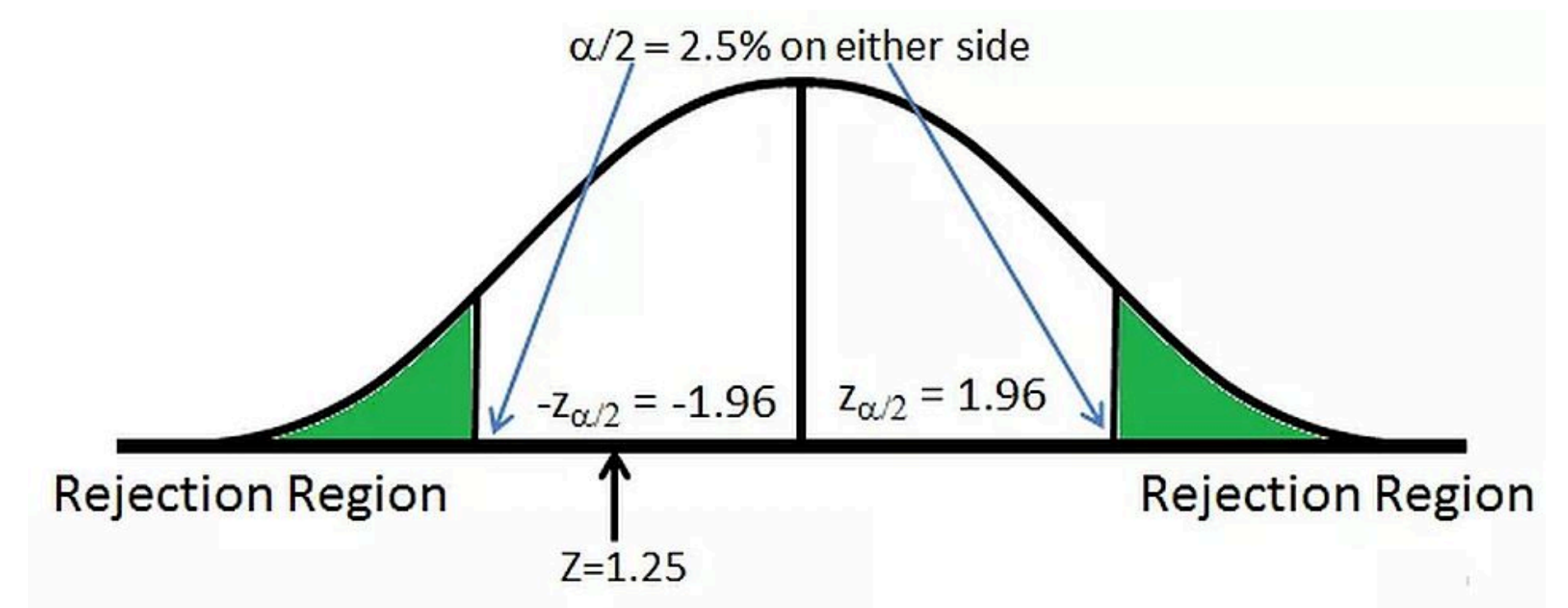
- Can construct sampling distribution assuming null hypothesis is true

- Set a **significance level** α for the test

- Fraction of distribution in each “tail” considered anomalous is $\alpha/2$ (if **two-sided test**)

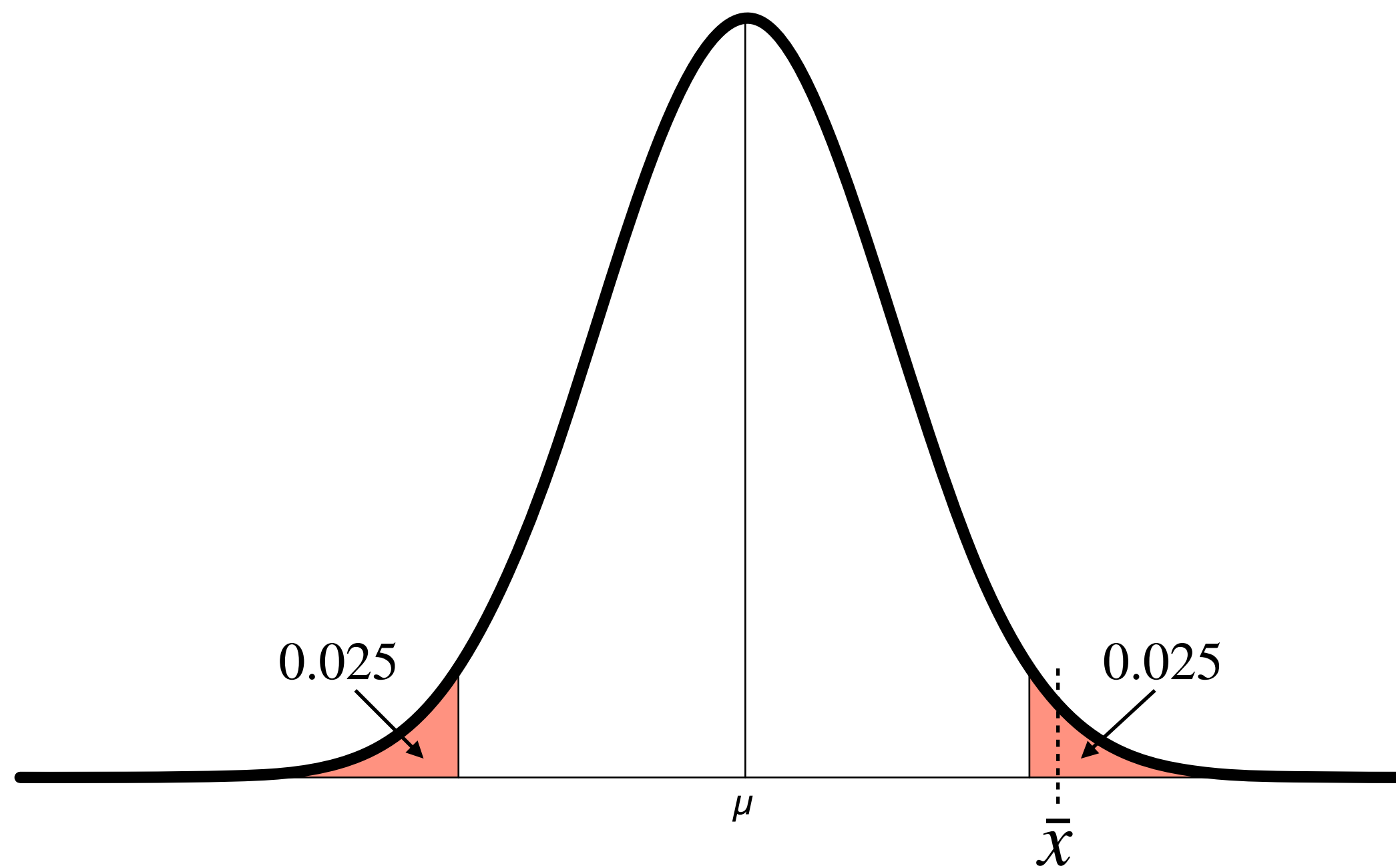
- See whether sample \bar{x} falls in that tail

- If so, *reject* null hypothesis H_0 in favor of alternative H_1 ; otherwise, *do not reject* (but this does not prove that H_0 is true)



z-test

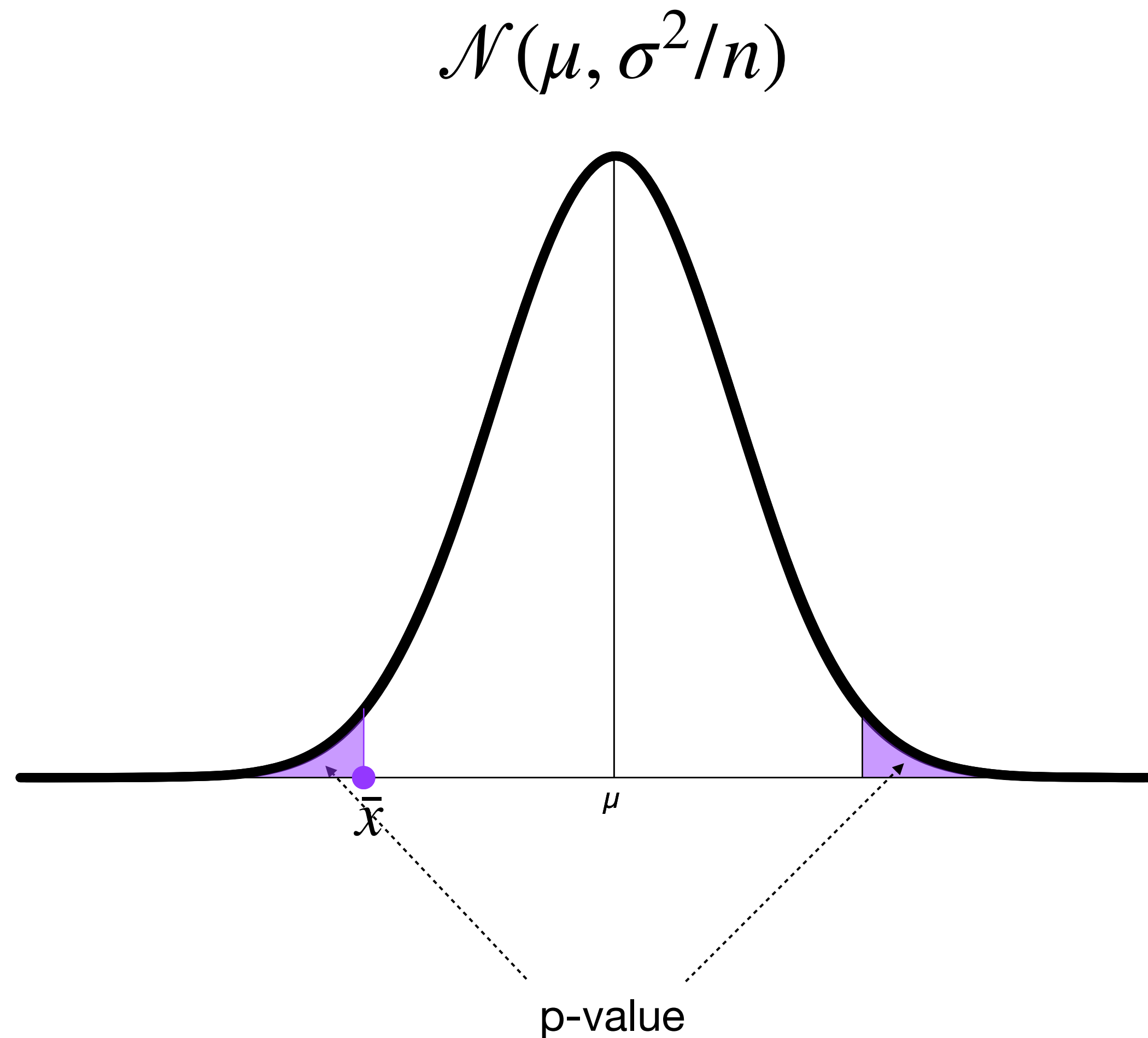
Significance level: $\alpha = 0.05$



In this case, reject H_0

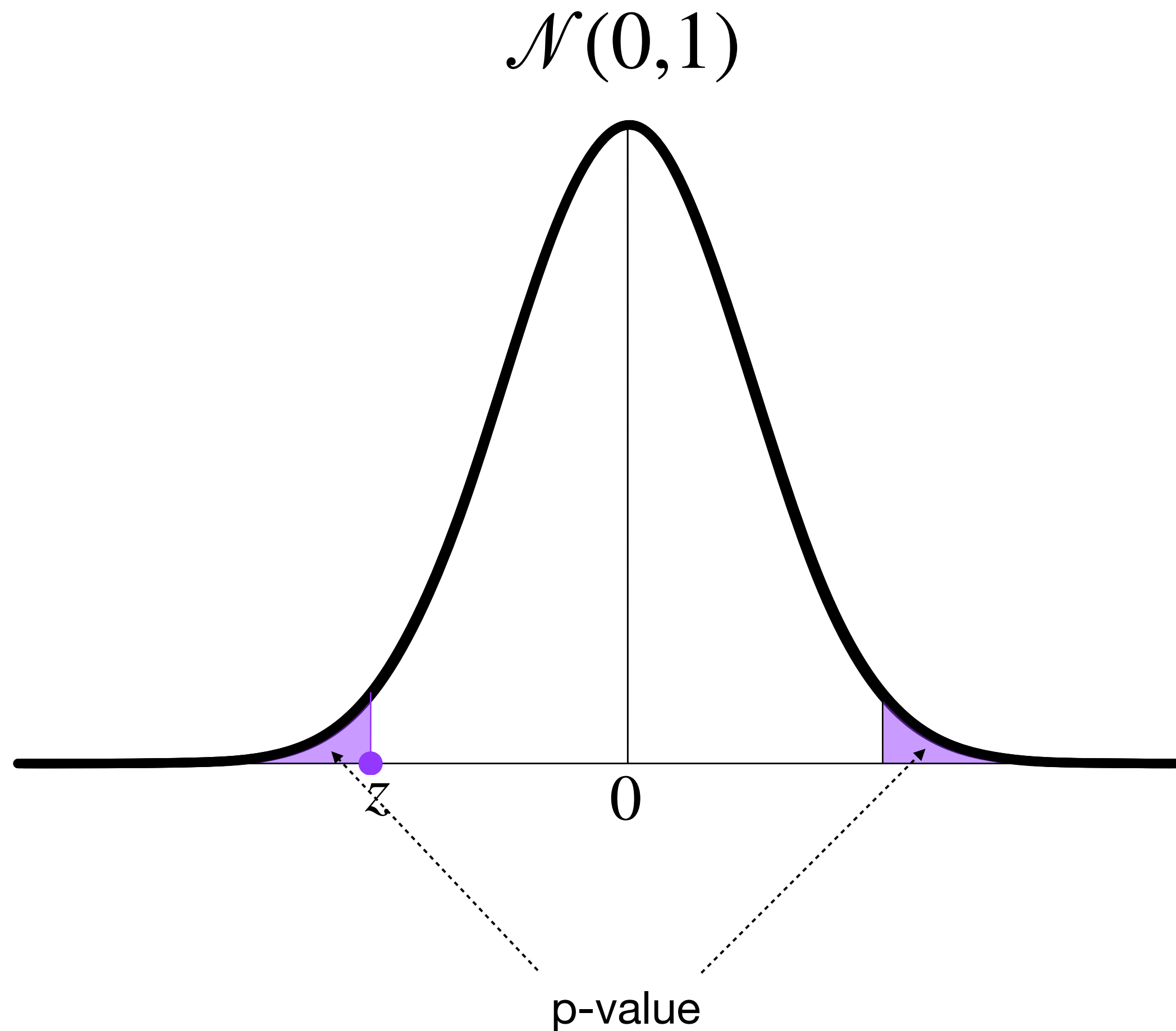
- Set a **significance level** α for the test
 - Fraction of distribution considered anomalous is $\alpha/2$ in each “tail” (if two-sided)
- See whether sample \bar{x} falls in that tail
- If so, *reject* null hypothesis H_0 in favor of alternative H_1 ; otherwise, *do not reject* (but this does not prove that H_0 is true)

p-value for z-test



- We can formalize this logic by calculating the **p-value**
- Place sample \bar{x} on distribution
- Ask what fraction of distribution is farther from the mean μ than the sample \bar{x}
- This is your p-value, which is compared to the significance level α :
 - Usually ask for $\alpha = 0.05$ or 0.01 (i.e., so that $p \leq 0.05, 0.01$ for significance)
 - Sometimes $\alpha = 0.1$ is OK

procedure for z-test



- Compute sample mean \bar{x}
- Compute standard deviation of sampling distribution (standard error)

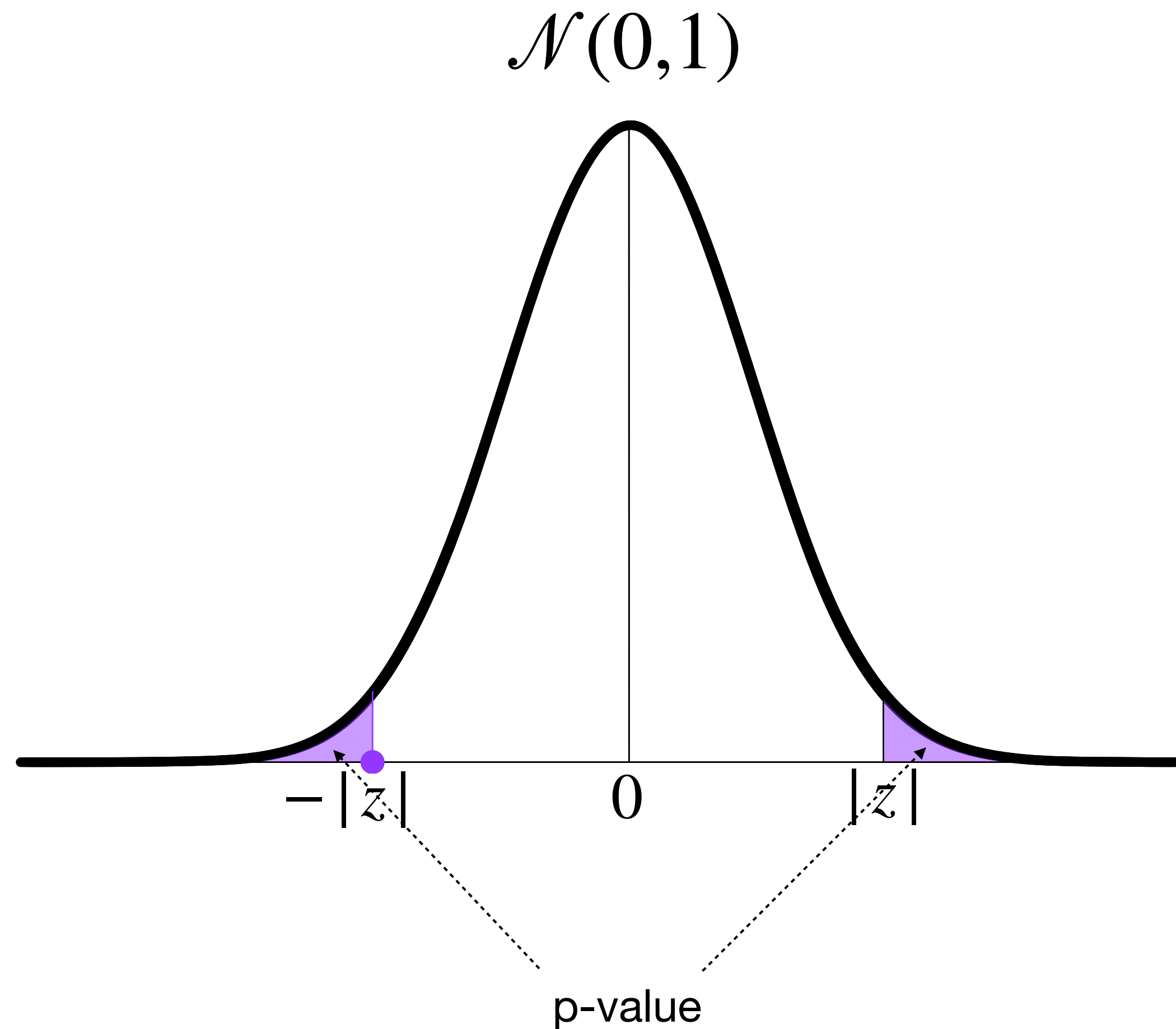
$$SE = \frac{\sigma}{\sqrt{n}}$$

- Compute **z-score**

$$z = \frac{\bar{x} - \mu}{SE}$$

- Normalizing the sample to the **standard normal distribution** $\mathcal{N}(0,1)$
- Compute p-value from z-score

computing p-value from z-score



- One way: look up in a standard table

- In Python:

```
import scipy.stats as stats
```

```
# compute  $z = (x - \mu) / SE$ 
```

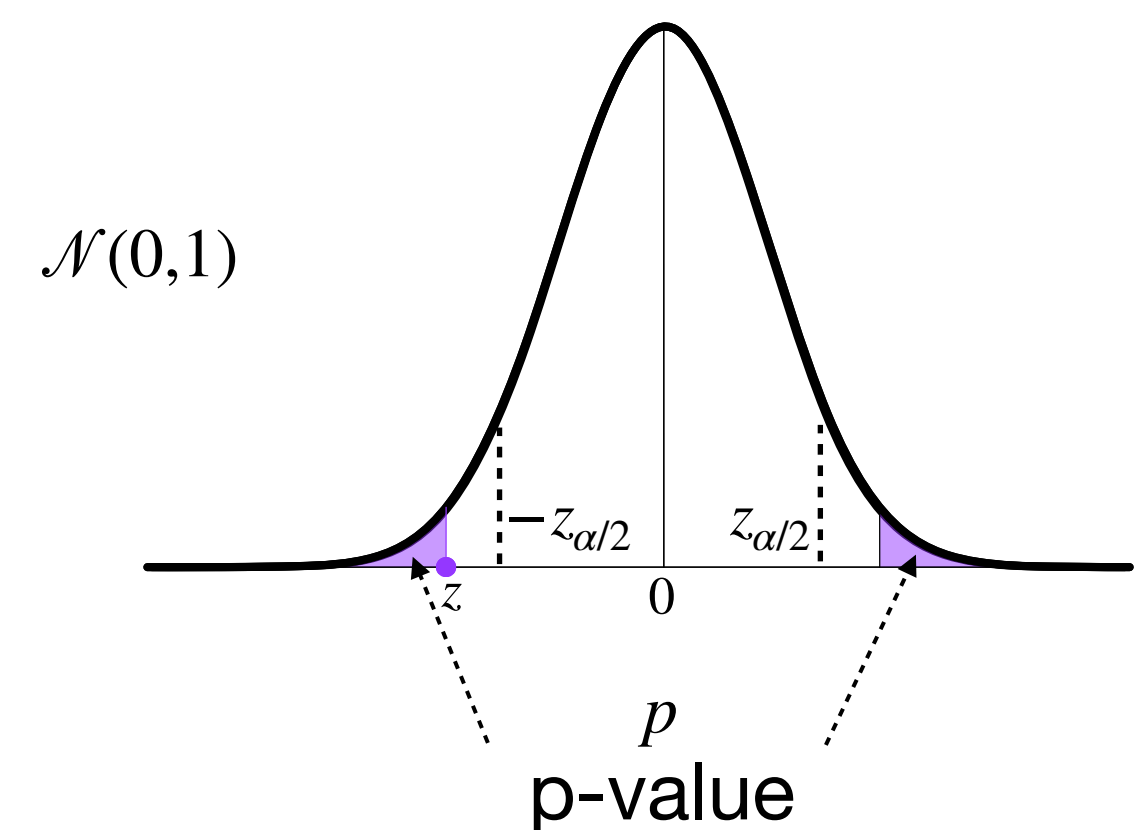
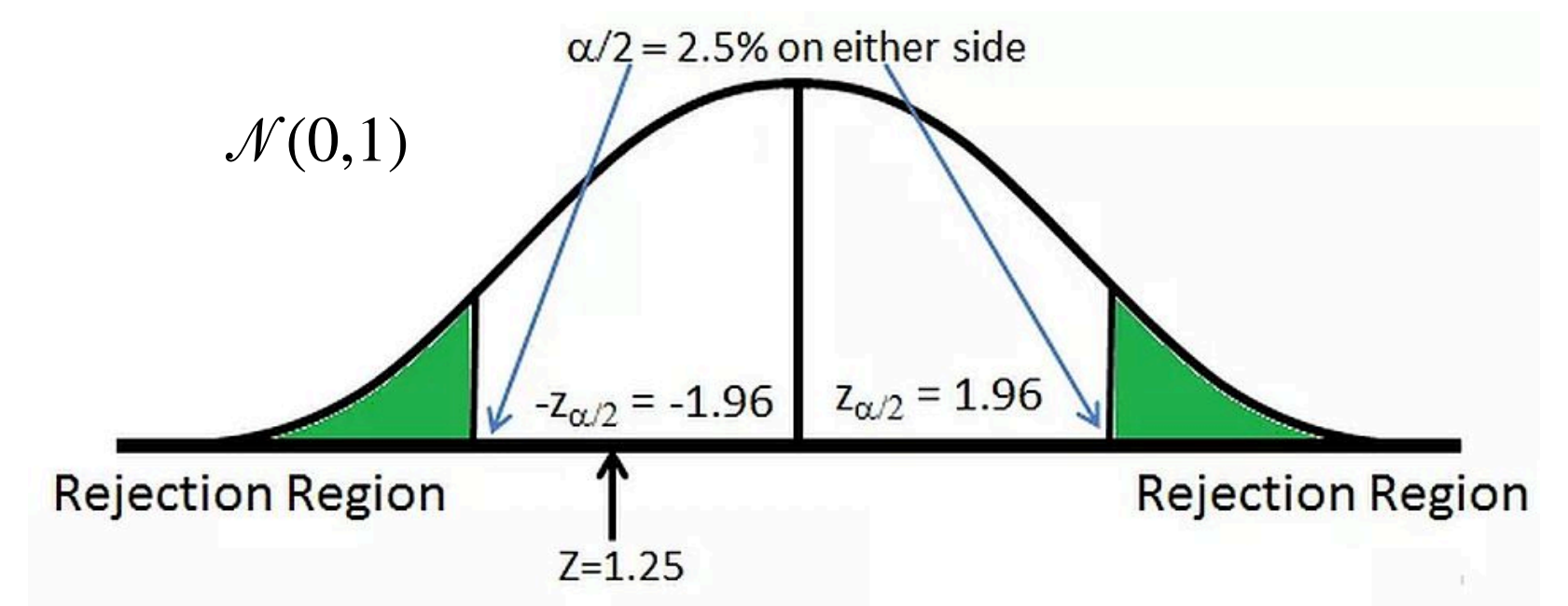
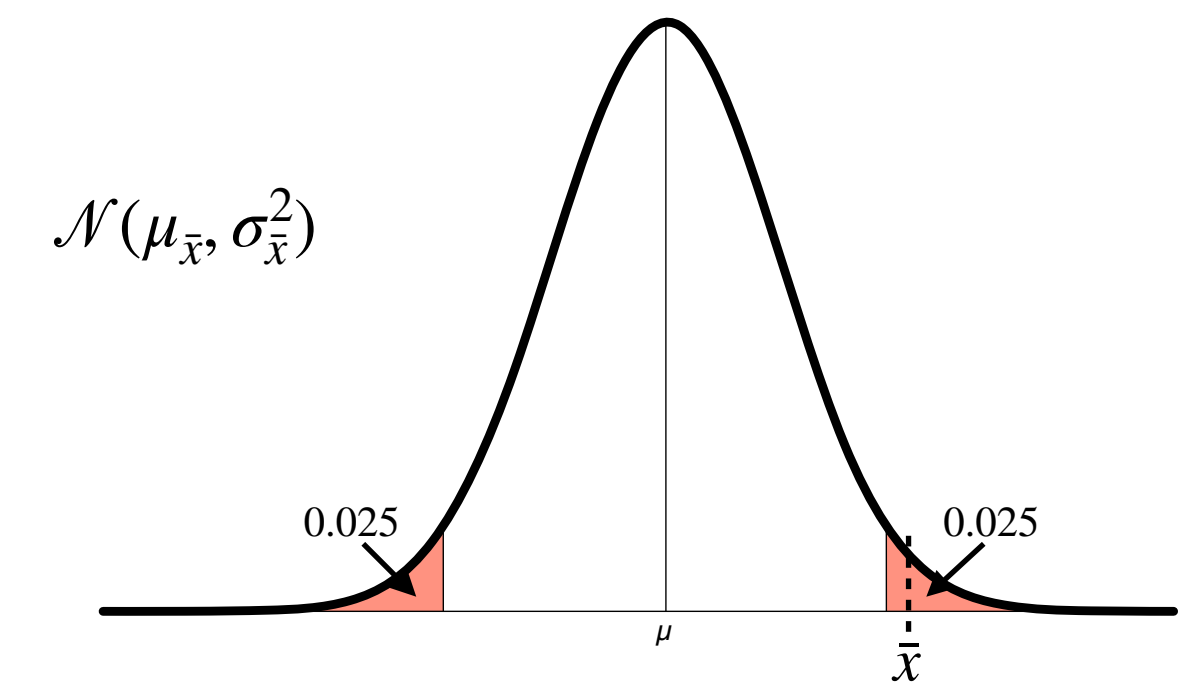
```
p = 2 * stats.norm.cdf(-abs(z))
```

- `norm.cdf()` is the CDF of $N(0,1)$, which we often denote $F_{N(0,1)}$
- Why `-abs(z)`? cdf considers left of the z point, so if z is positive, we want to reference $-z$

overview of z-test

- Assumptions needed for statistical test
 - Null hypothesis H_0
 - Alternative hypothesis H_1
 - A statistical significance level α
- Equivalent questions (if yes, then reject null hypothesis)
 - Is the sample mean, \bar{x} , in tail defined by α of the **sampling distribution** $\approx \mathcal{N}(\mu, \sigma_{\bar{x}}^2)$?
 - Is the z-score, $z = \frac{\bar{x} - \mu_{\bar{x}}}{\sigma_{\bar{x}}} = \frac{\bar{x} - \mu}{SE}$, in the tail defined by α of a **standard normal** $\mathcal{N}(0,1)$?
 - Is the **p-value**, $p = 2F_{\mathcal{N}(0,1)}(-|z|)$, less than α ?

Significance level: $\alpha = 0.05$



back to our original example

- $\mu = 100, \sigma = 22$

$$\bar{x} = 95, n = 100$$

- So we calculate:

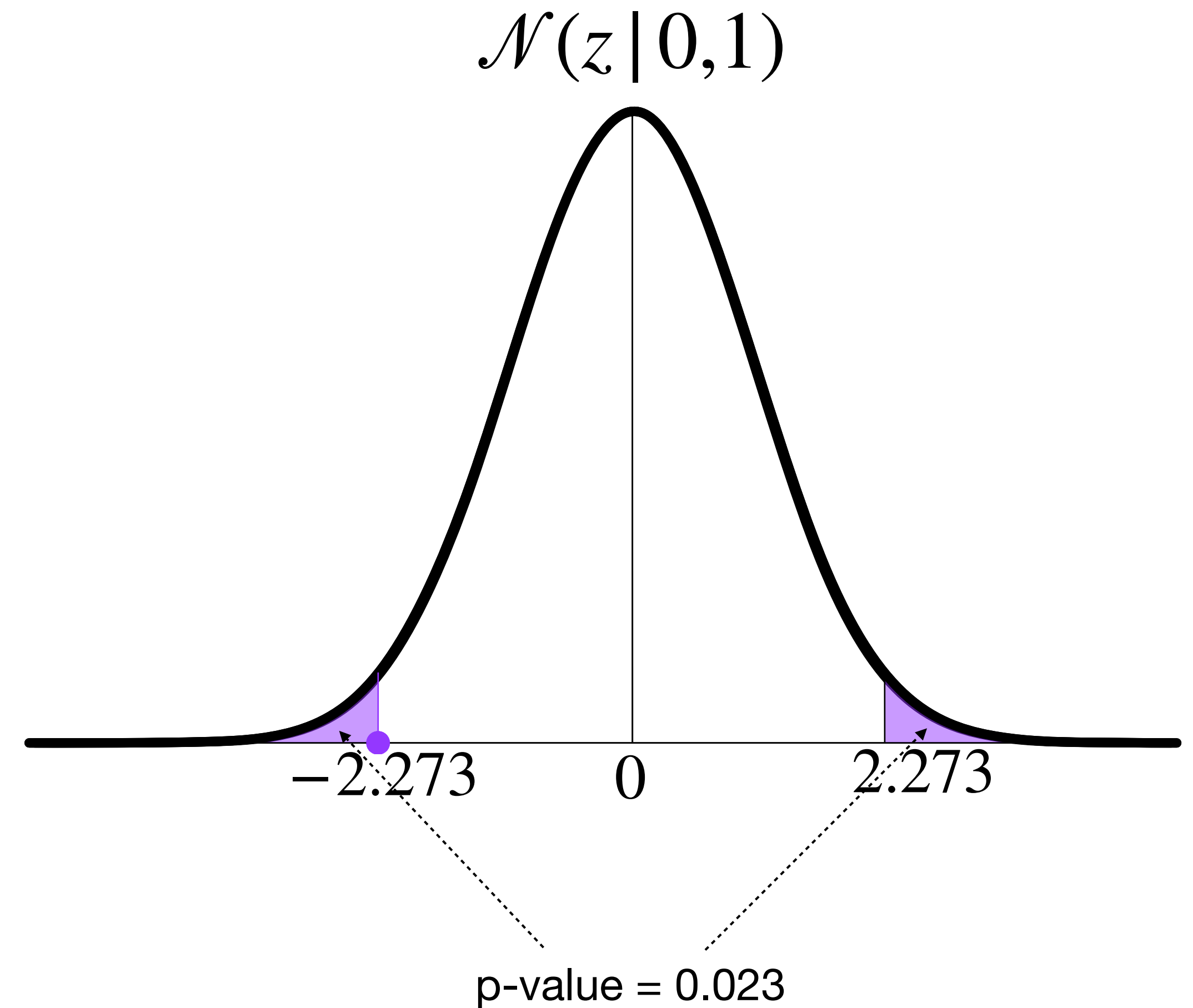
$$z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} = \frac{95 - 100}{22/\sqrt{100}} = -2.273$$

$$p = 2 \cdot F(z | 0,1) = 0.023$$

- Conclusion:

- Significant at $\alpha = 0.1, 0.05$ (reject H_0)

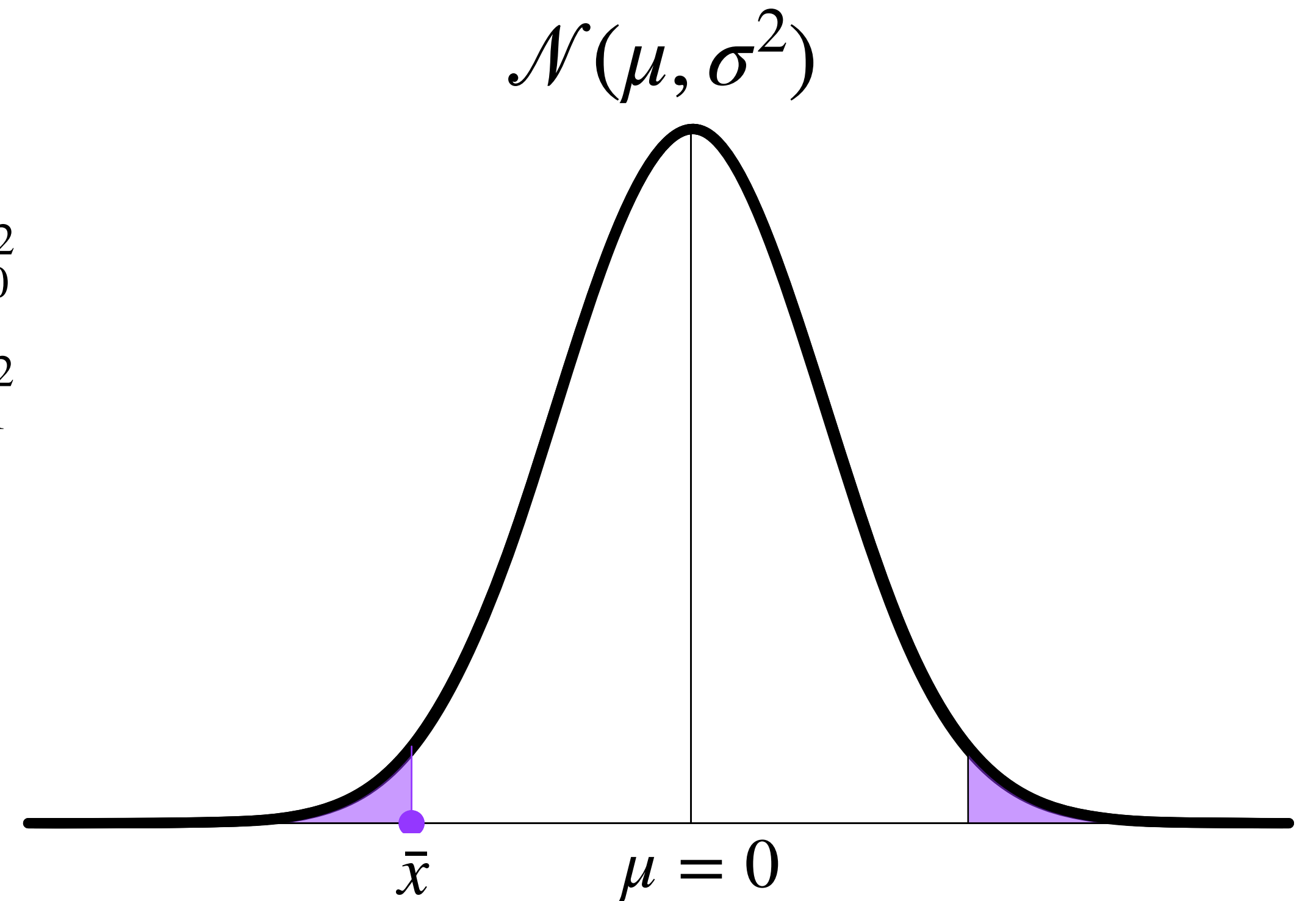
- Not significant at $\alpha = 0.01$ (cannot reject H_0)



comparing two means

- What if you have *two* populations, and you want to know whether their means are statistically different?
 - Sample 1: Sample size n_0 , from pop. mean μ_0 , variance σ_0^2
 - Sample 2: Sample size n_1 , from pop. mean μ_1 , variance σ_1^2
- Hypotheses
 - H_0 : The means are the same, i.e., $\mu_0 = \mu_1$
 - H_1 : The means are different, i.e., $\mu_0 \neq \mu_1$
- Can use **two-sample z-test**
- Under null hypothesis, sampling distribution of *difference between two means* has:

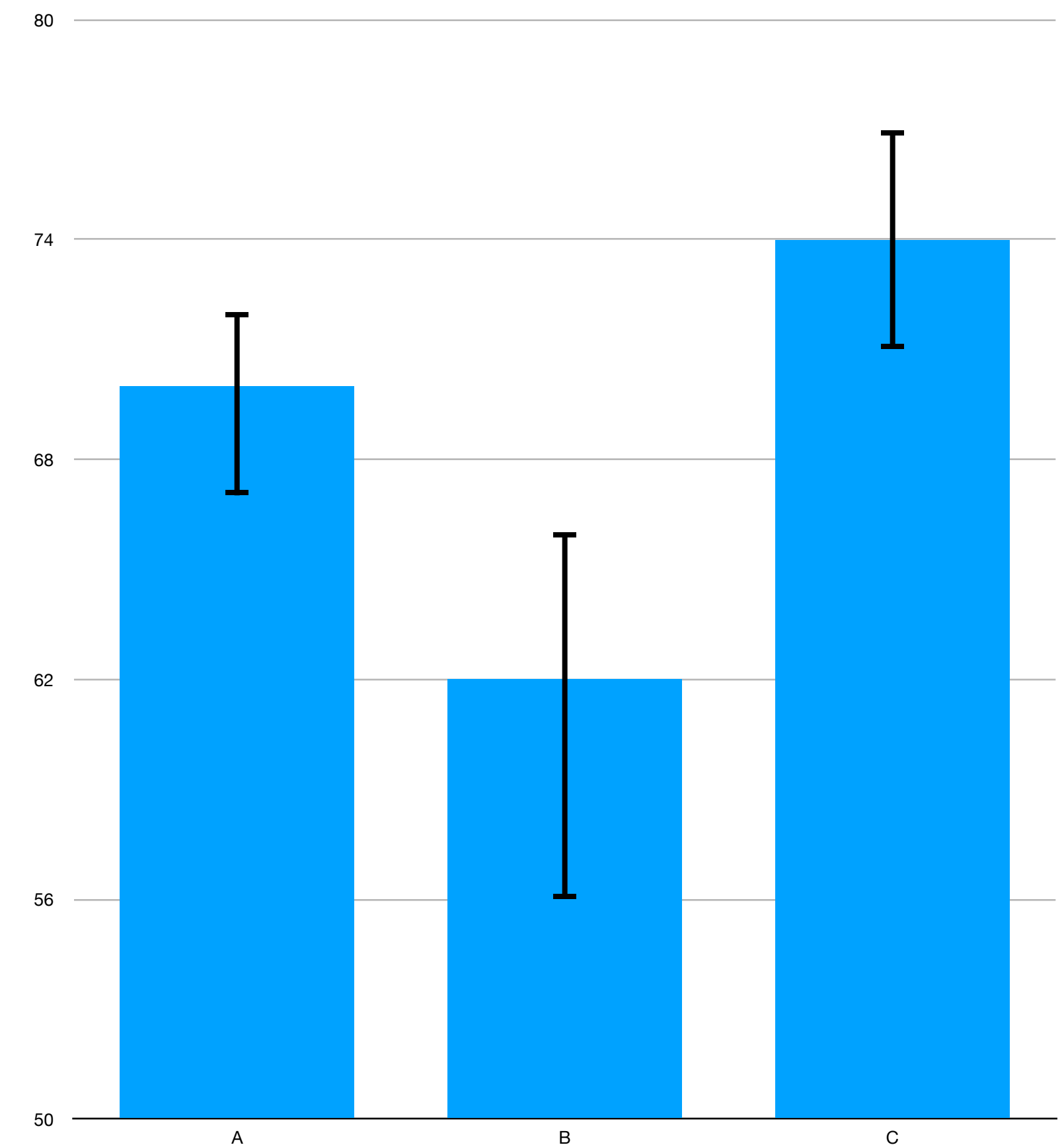
$$\mu = \mu_0 - \mu_1 = 0 \qquad \sigma = \sqrt{\frac{\sigma_0^2}{n_0} + \frac{\sigma_1^2}{n_1}}$$



- Test point is $\bar{x} = \bar{x}_0 - \bar{x}_1$
- z-score is $(\bar{x} - \mu)/\sigma$

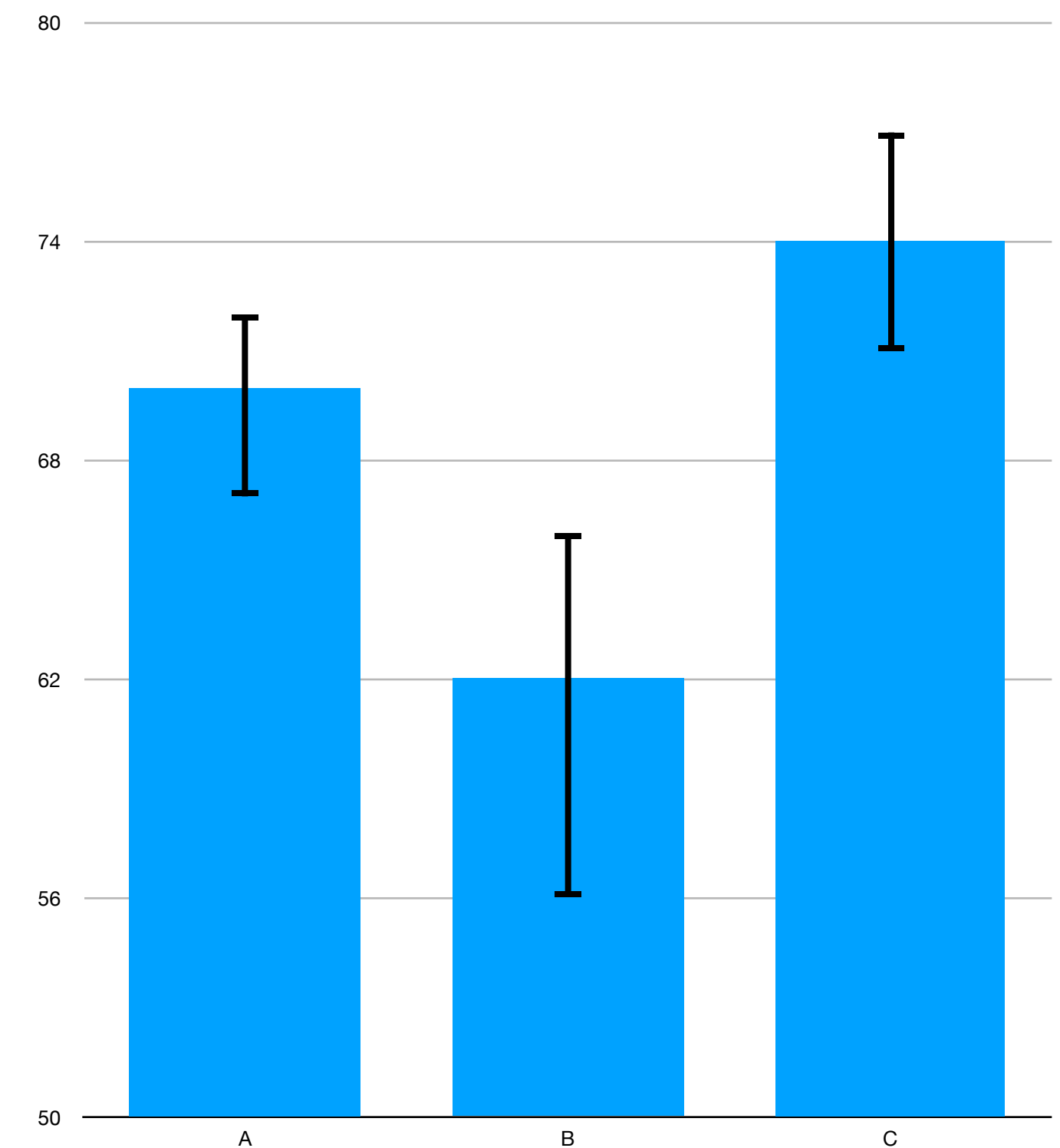
confidence intervals

- We see these a lot: Ranges above and below values on a graph
 - What do they mean?
- Surprisingly tricky question to answer



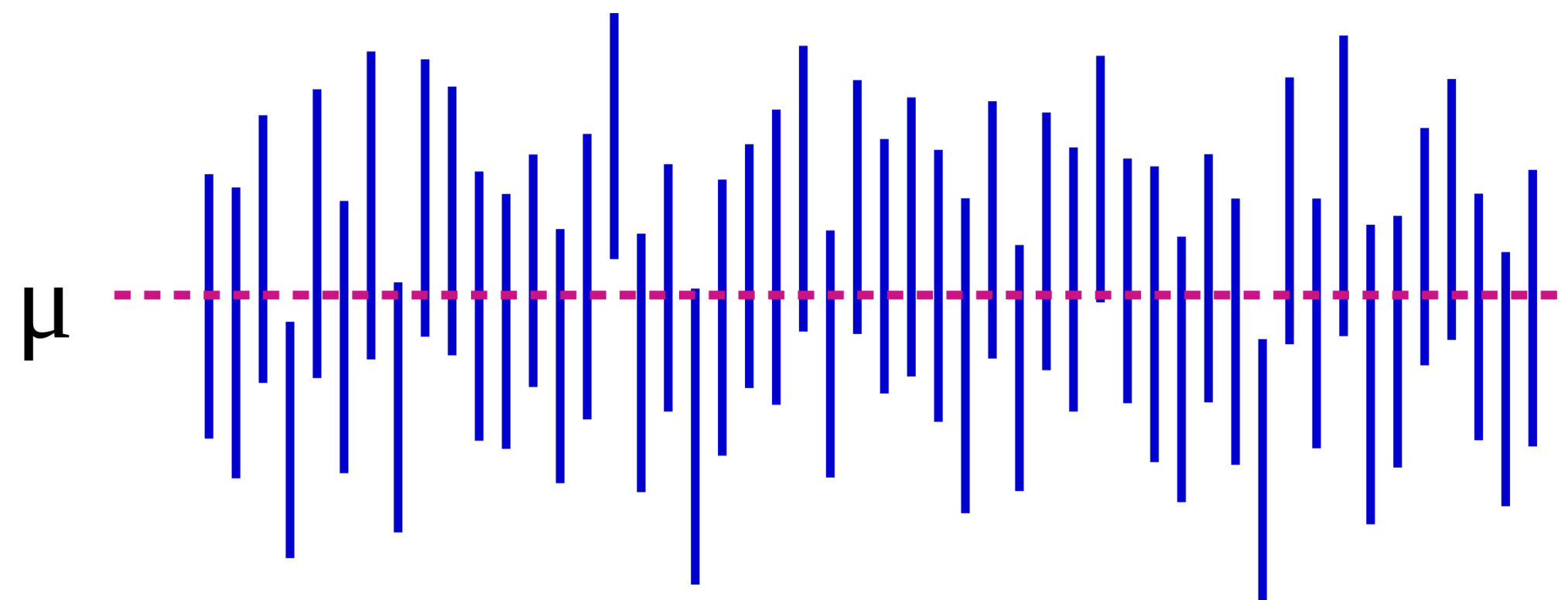
intuition of confidence intervals

- A **confidence interval** is a range around the mean which says something about how “good” your estimation procedure is
 - How “good” is your choice of number of samples, given the variance in the population
- Interpretation of a (95%) confidence interval:
 - *if I were to repeat the experiment a large number of times, 95 percent of confidence intervals would contain the population mean*
 - *before I run the experiment, there is a 95 percent chance that the population mean will fall within the computed confidence interval*
 - *if the population mean is inside the confidence interval, it would not be statistically significant (informally, you wouldn't be surprised!)*



the first interpretation

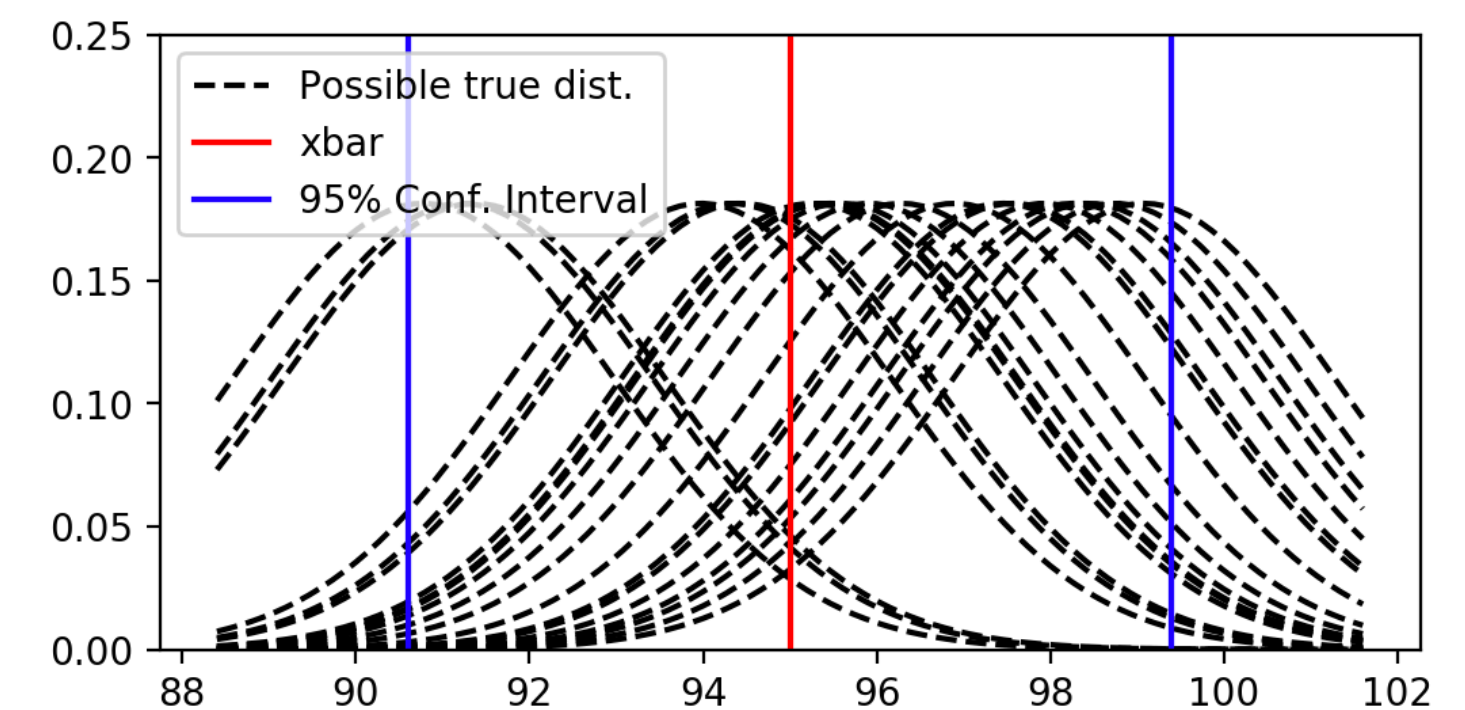
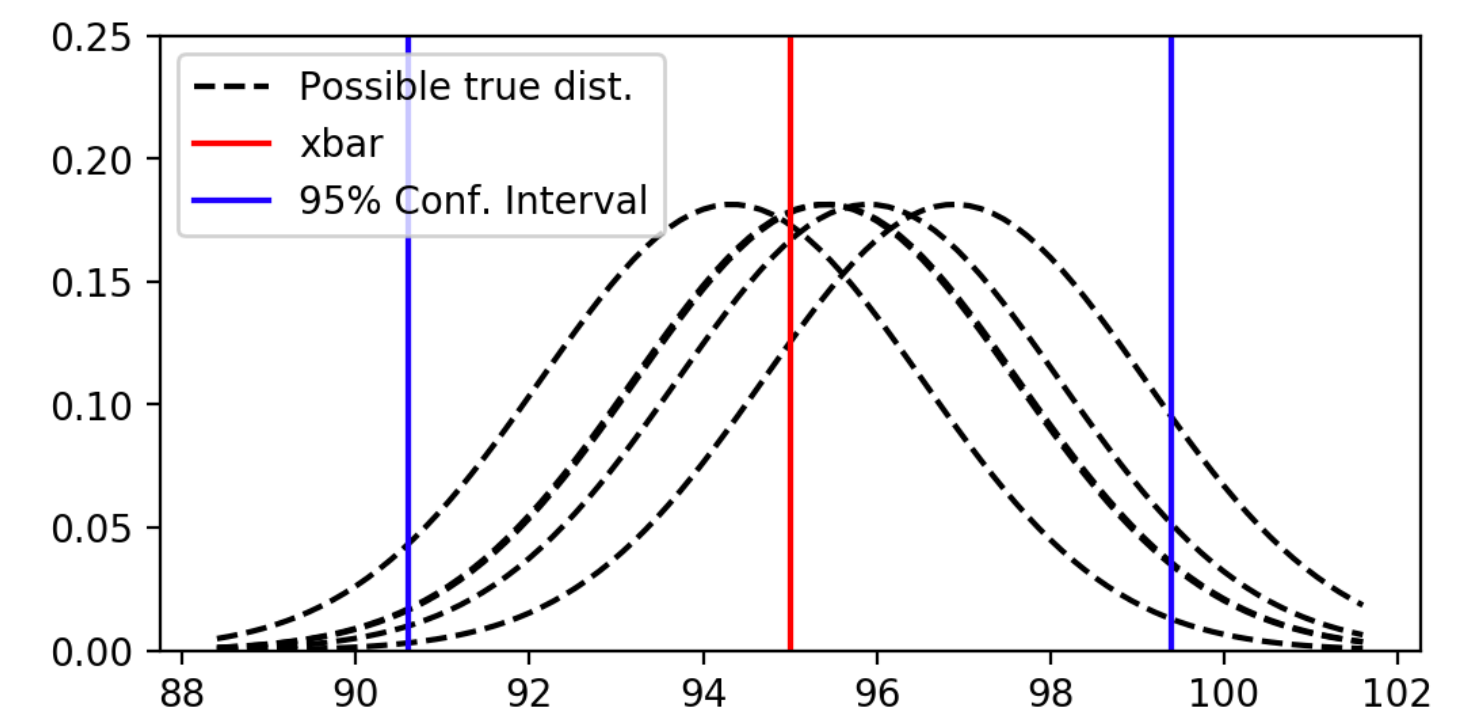
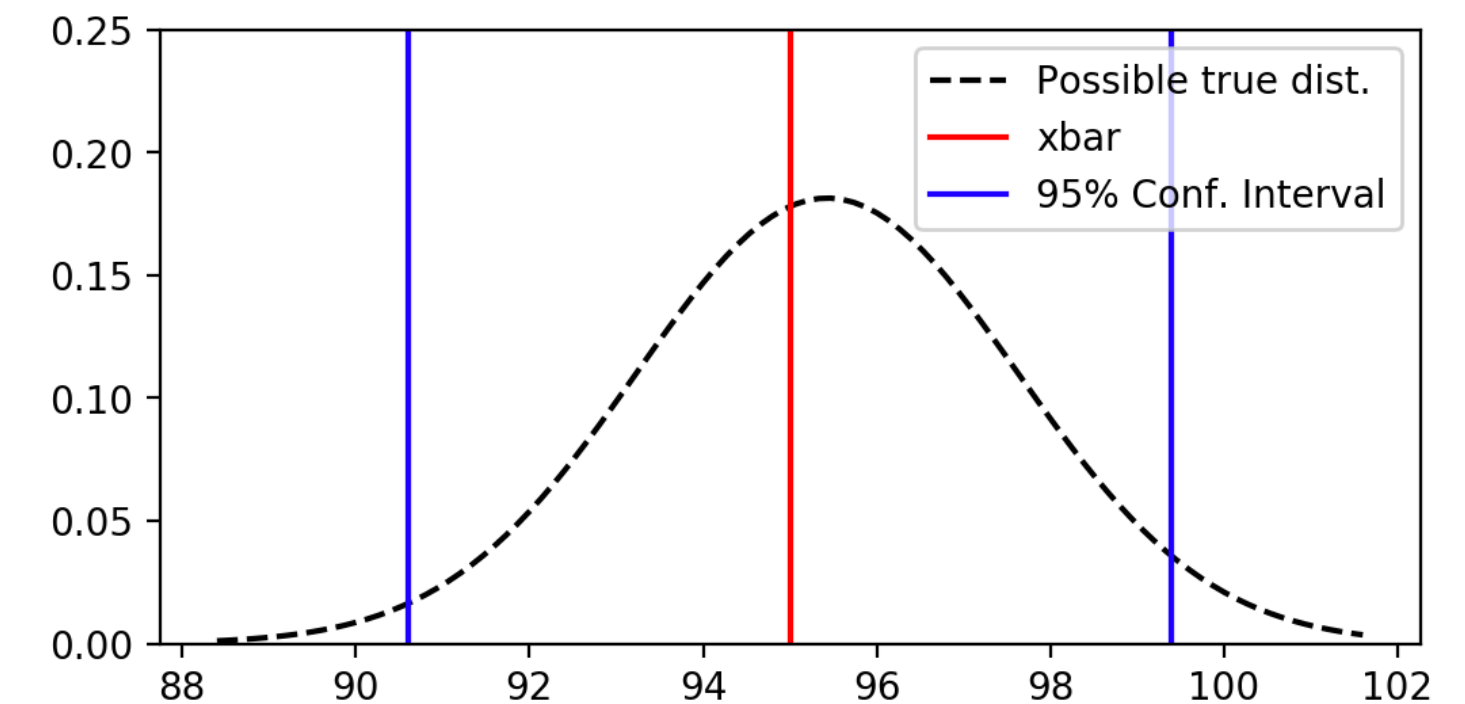
- *If I were to repeat the experiment a large number of times, 95 percent of confidence intervals would contain the population mean*
- In the diagram below, each vertical bar is one confidence interval calculated for one experiment
- For a 95% confidence interval, we expect 95% of them will include μ



source: NYW-confidence-interval.svg
Wikipedia user Tsyplakov

confidence intervals more formally

- If the population parameter is outside the $c\%$ **confidence interval**, then an event occurred that had a probability of less than $(100 - c)\%$ of happening
- Note that we are setting c ahead of time (unlike with hypothesis testing, where we figure out how likely/unlikely something is *after* the fact)
 - *Wide confidence interval*: The variance of your data is high (and/or your sample size is small), so we need a wide interval to make the above statement true.
 - *Narrow confidence interval*: The variance of your data is small (and/or your sample size is large), so we *don't* need a wide interval to make the above statement true.



computing confidence intervals

- Conceptually related to z-tests, but the perspective is *reversed*
 - For what sampling distributions (centered at the population mean), would our sample mean NOT be surprising?
 - Note: Our confidence interval is centered around the *sample mean* (instead of the hypothesized population mean)
- Remember definition of z-score:

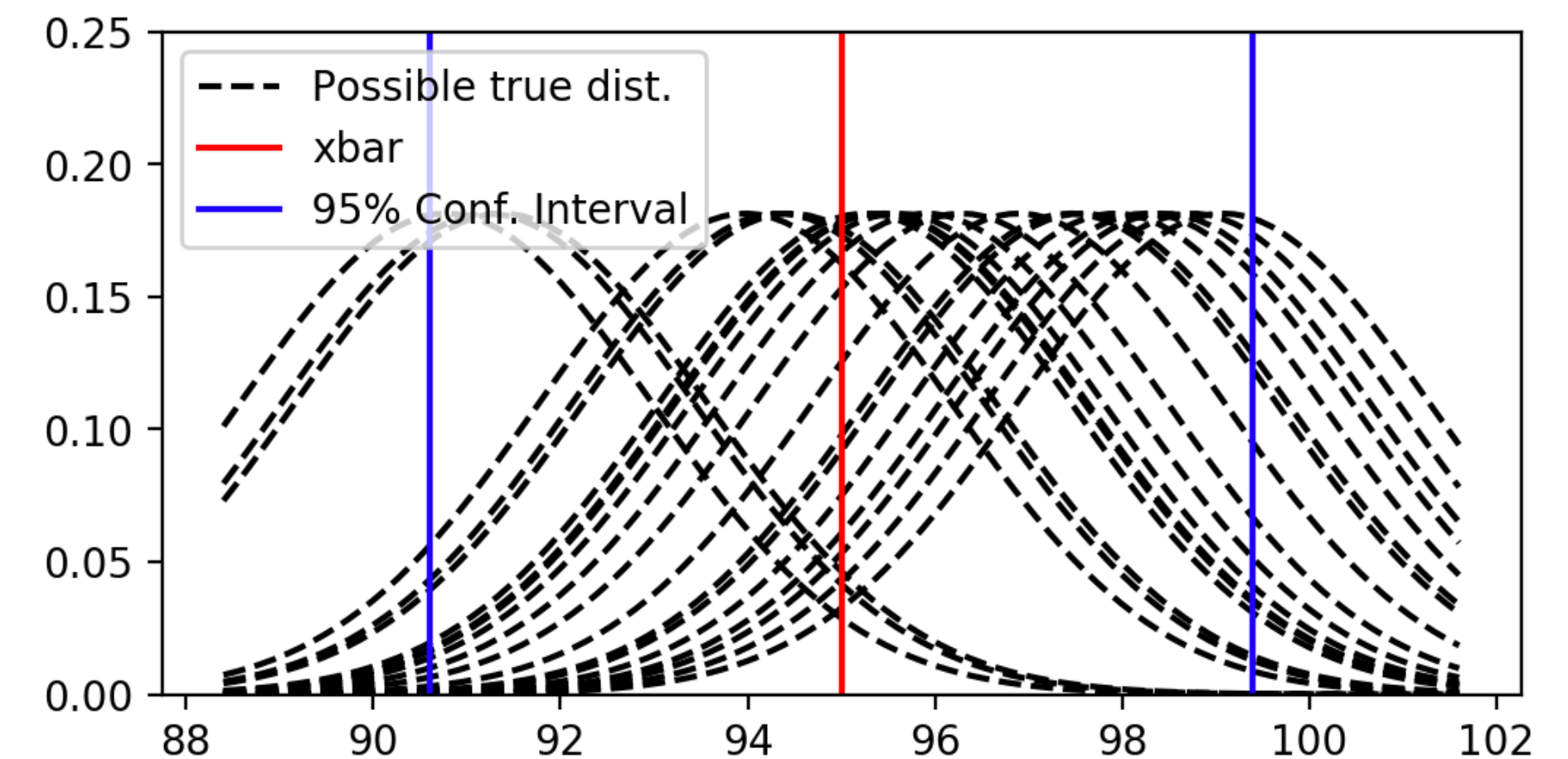
$$z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}}$$

- And p-value:

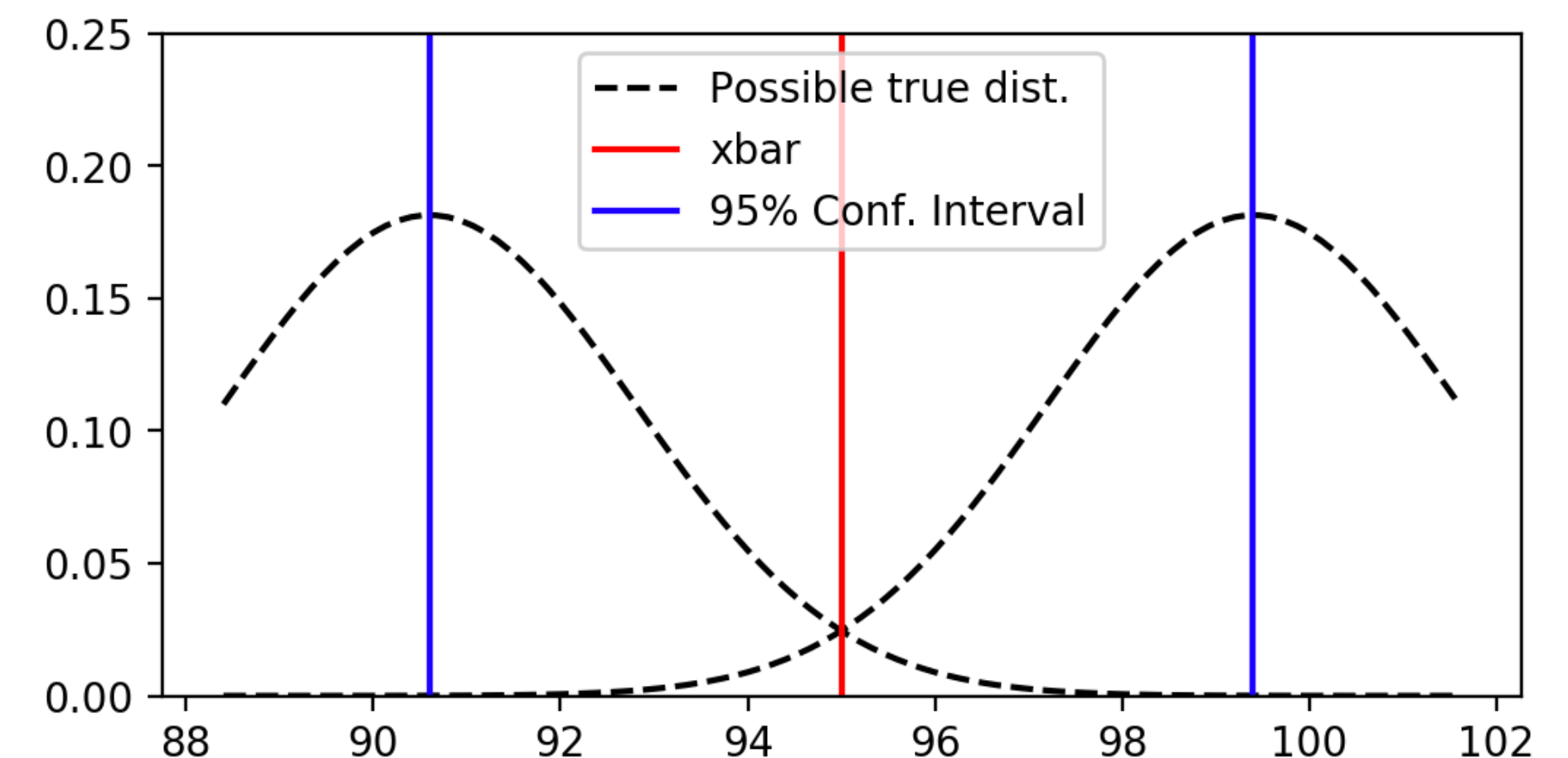
$$p = 2 * \text{sp.stats.norm.cdf}(-\text{abs}(z))$$

- If c is the desired confidence level (here in decimal form), what z do we need such that $p \leq (1 - c)$?

Possible values of μ such that \bar{x} would be unsurprising



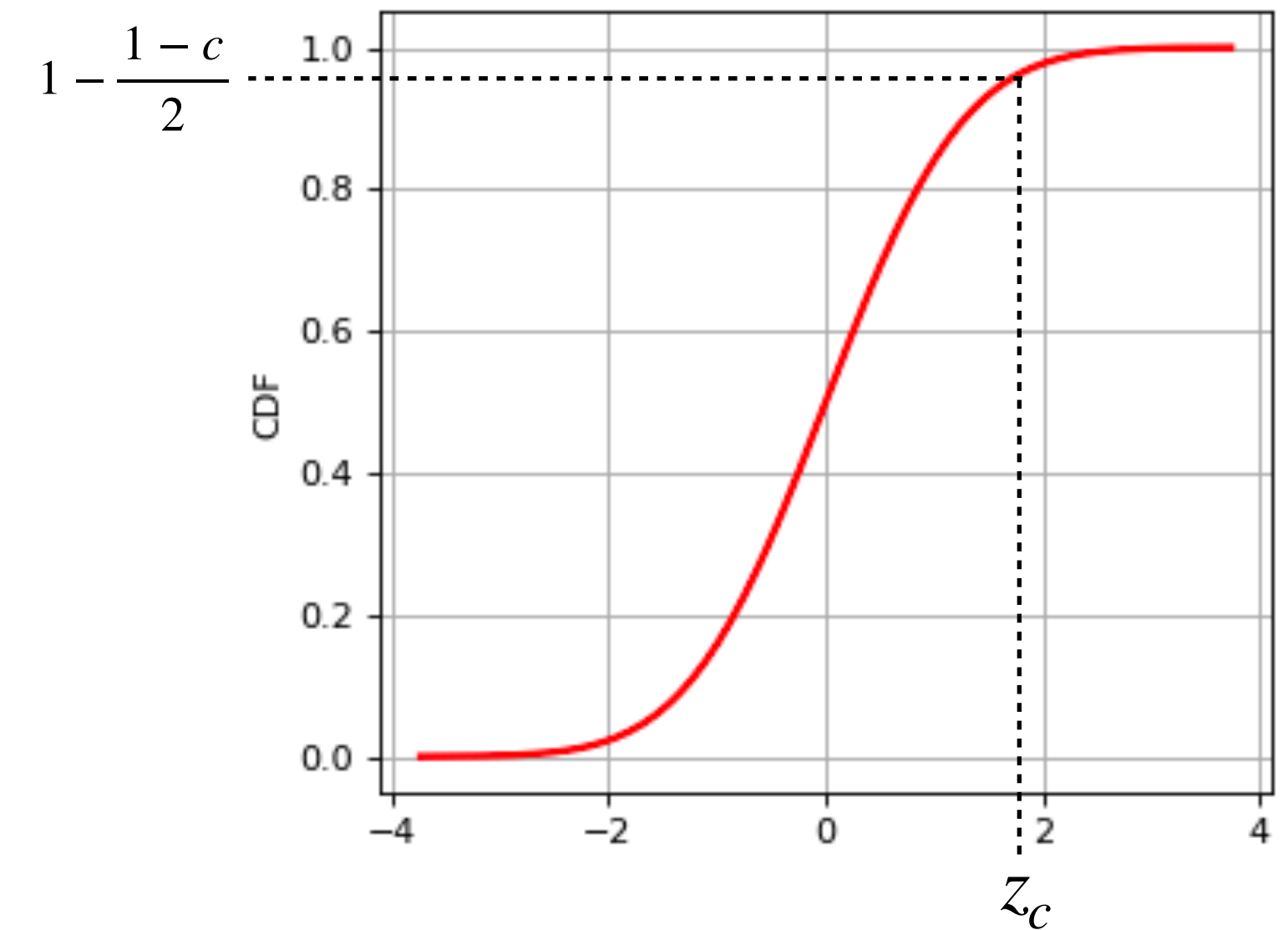
Extreme values of μ such that \bar{x} is still unsurprising



computing confidence intervals

- Call this z_c
- Compute in Python as follows:

```
z_c = stats.norm.ppf(1 - (1 - c)/2)
```
- While `norm.cdf` goes from z-score to probability, `norm.ppf` goes from probability to z-score

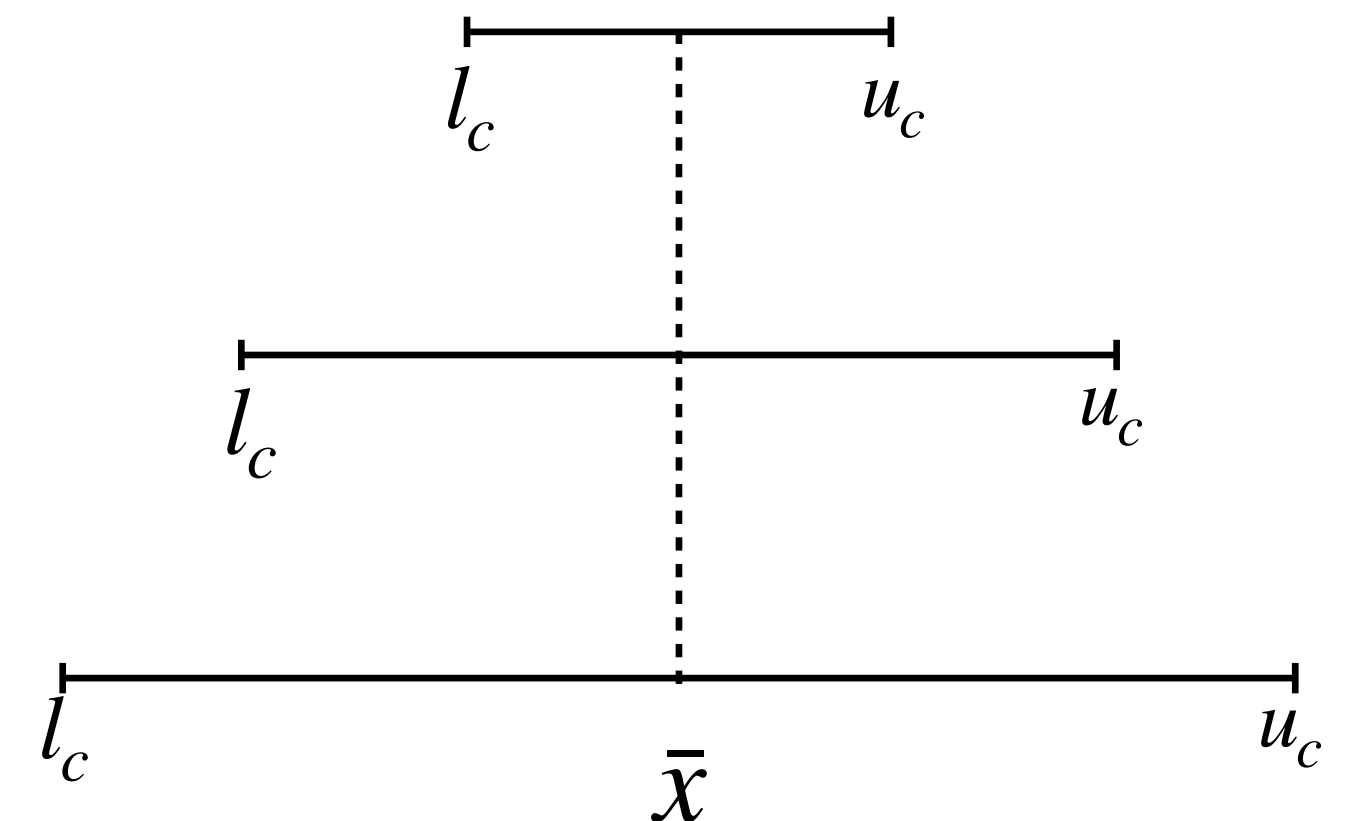


- Now we can answer the question: *What range of μ would be “unsurprising” at $c\%$ confidence level?*

$$z_c = \left| \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} \right| \rightarrow \mu \in \left(\bar{x} - \frac{z_c \cdot \sigma}{\sqrt{n}}, \bar{x} + \frac{z_c \cdot \sigma}{\sqrt{n}} \right) = (l_c, u_c)$$

- This is your $c\%$ confidence interval

increasing
 c



back to our original example

- Let's calculate 90%, 95%, and 99% confidence intervals for μ

- Recall that our sample had

$$\bar{x} = 95g, \sigma = 22g, n = 100$$

- Thus, the confidence intervals are:

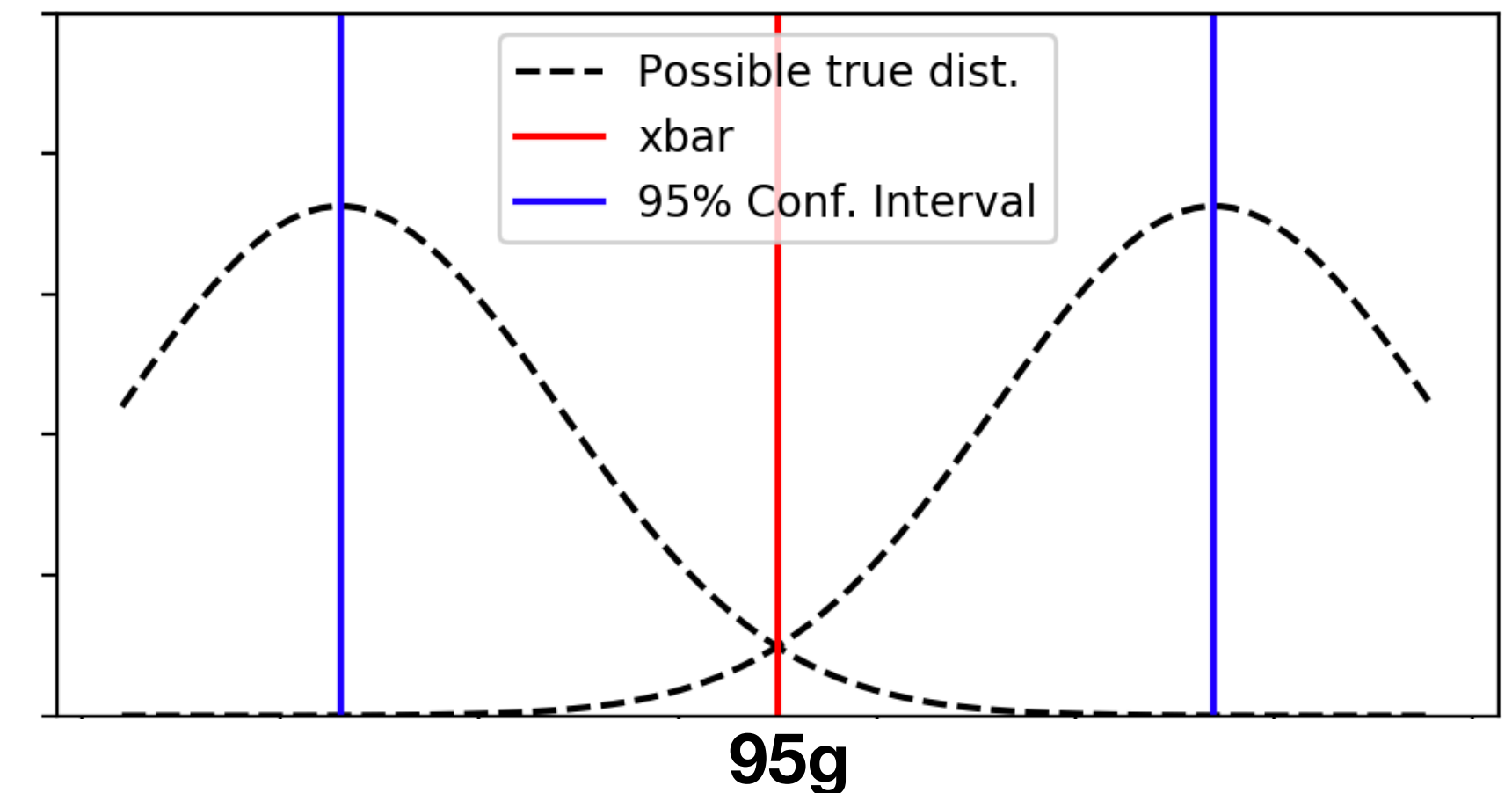
$$\mu \in \left(95 - \frac{\sigma}{\sqrt{n}} \cdot z_c, 95 + \frac{\sigma}{\sqrt{n}} \cdot z_c \right) = (95 - 2.2 \cdot z_c, 95 + 2.2 \cdot z_c)$$

- For 90%, 95%, 99%, $z_c = 1.645, 1.960, 2.576$. Thus,

$$90 \% : (91.38, 98.62)$$

$$95 \% : (90.69, 99.31)$$

$$99 \% : (89.33, 100.67)$$

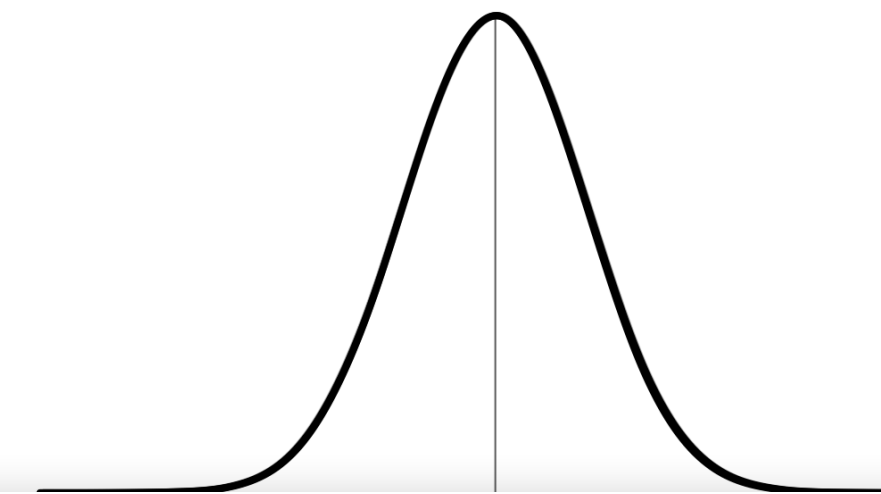


How would we make the intervals narrower for the same levels of confidence?

we've been fudging

- Recall that to use the z -distribution, we must either know σ or have large enough n
- The **student's t-distribution** and **t-test** is used when the normal approximation does not hold:
 - i.e., when we don't know σ (which we usually do not) **and** when $n < 30$
 - Can use this to reason about μ , including building confidence intervals and conducting hypothesis tests

hypothesis testing



- Suppose the null hypothesis **was true** (new widgets are the same as the original widgets)
- Then the sampling distribution should have its mean at 100g
- And the sampling distribution should have a standard deviation of:

$$\sigma = \frac{22}{\sqrt{20}}$$

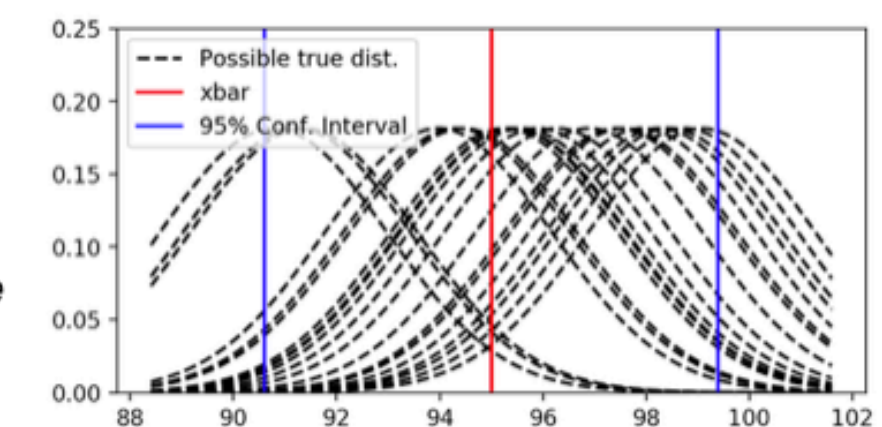
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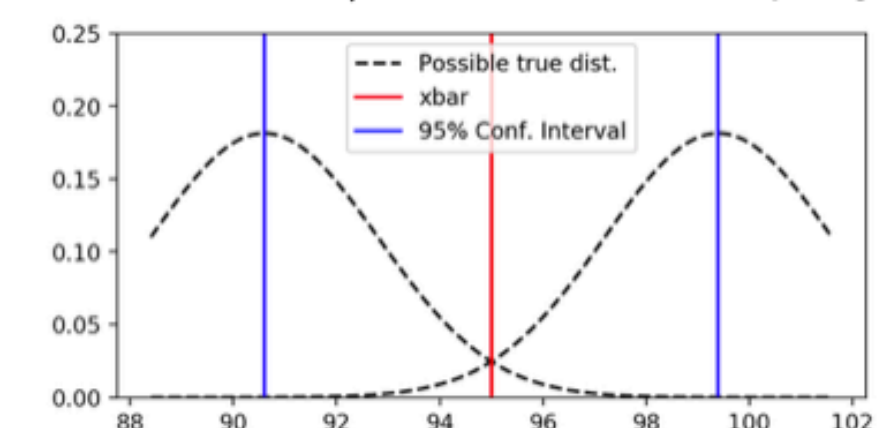
$$z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}}$$

- And p-value:
`p = 2 * sp.stats.norm.cdf(-abs(z))`
- If c is the desired confidence level (here in decimal form), what z do we need such that $p \leq (1 - c)$?

Possible values of μ such that \bar{x} would be unsurprising

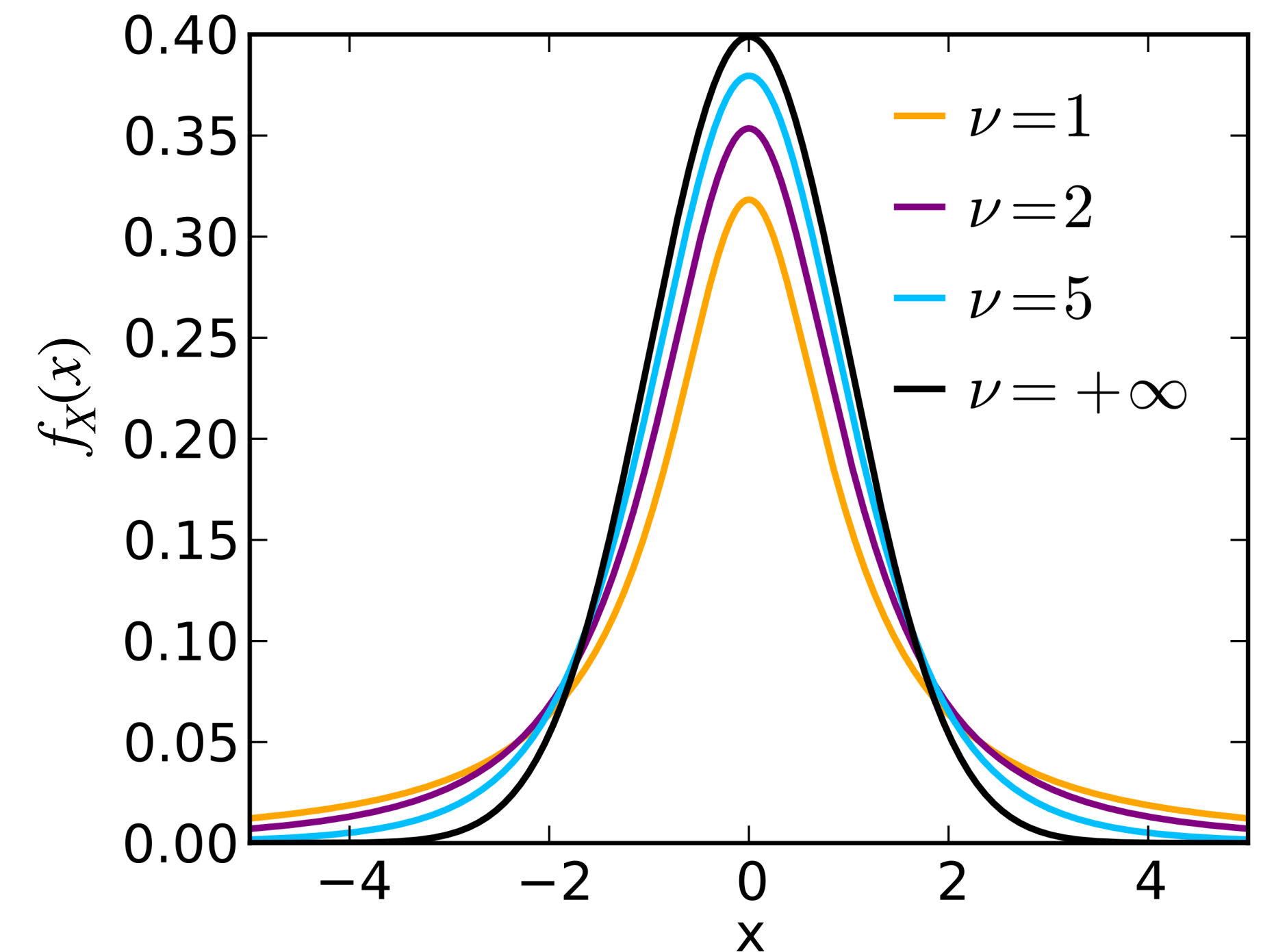


Extreme values of μ such that \bar{x} would be unsurprising



student's t-distribution

- Similar to the standard $\mathcal{N}(0,1)$ normal distribution (density shown to the right)
 - Symmetric about mean
 - Bell curve shaped
- But has **fatter tails**, i.e., more weight of the distribution away from the mean
 - Accounts for outliers better
- Parameter of the distribution is the **degrees of freedom ν**
 - $\nu = n - 1$: One less than the number of samples
 - Looks more and more like the standard normal as $n \rightarrow \infty$



t-test and confidence intervals

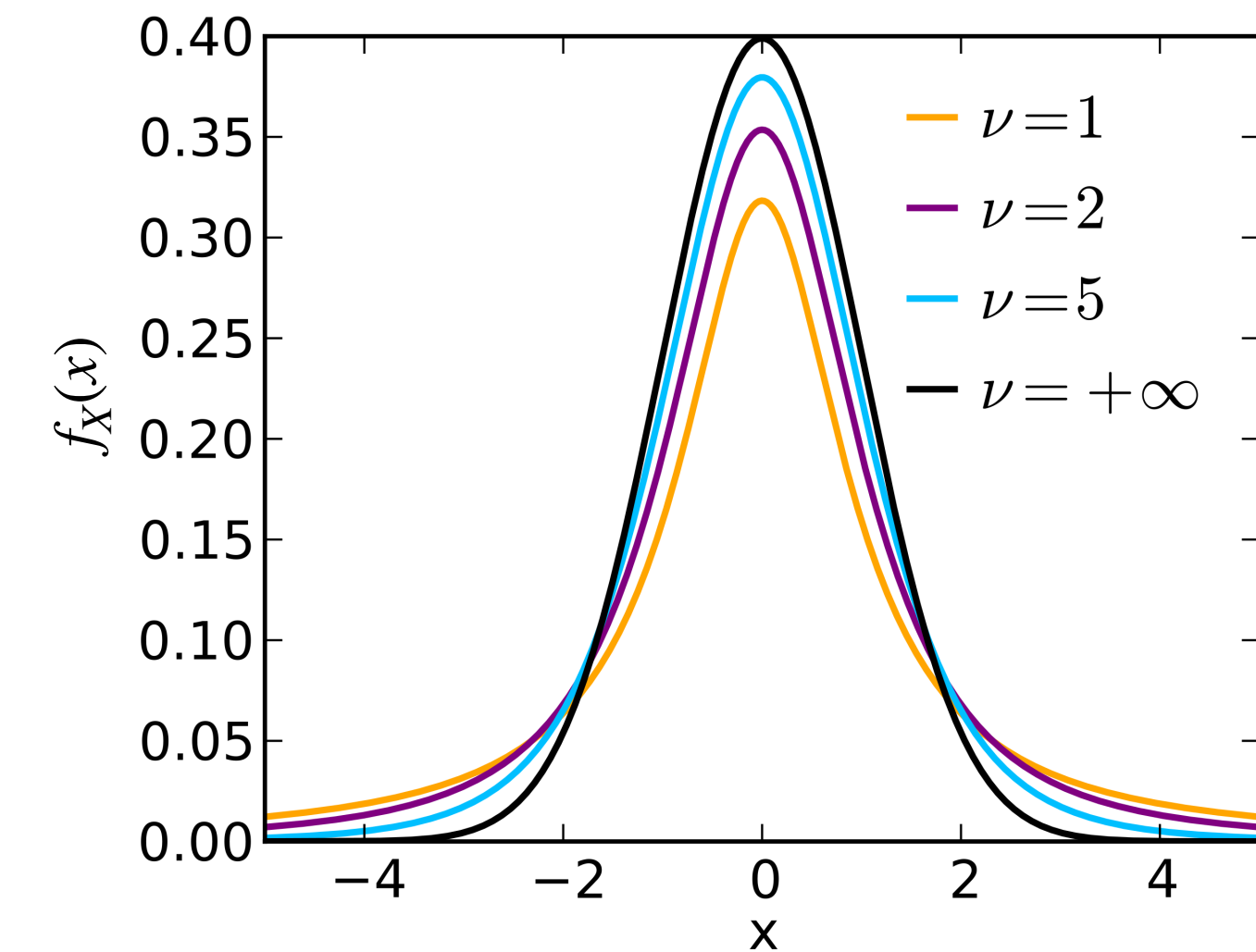
- Works the same as the z -test, except
 - Use s instead of σ
 - Compare to the t -distribution
- Computing the test statistic:
 - First get the standard deviation of the sample:

$$s = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2}$$

- Then we get the “ t -score”:

$$t = \frac{\bar{x} - \mu}{s/\sqrt{n}}$$

Compare to the
formula for z



- Then we get the p -value:
 $p = 2 * \text{stats.t.cdf}(-\text{abs}(t), \text{df})$
- And for confidence intervals, we find the t -score corresponding to c :

$$t_c = \text{stats.t.ppf}(1 - (1 - c)/2, \text{df})$$

one-sided tests

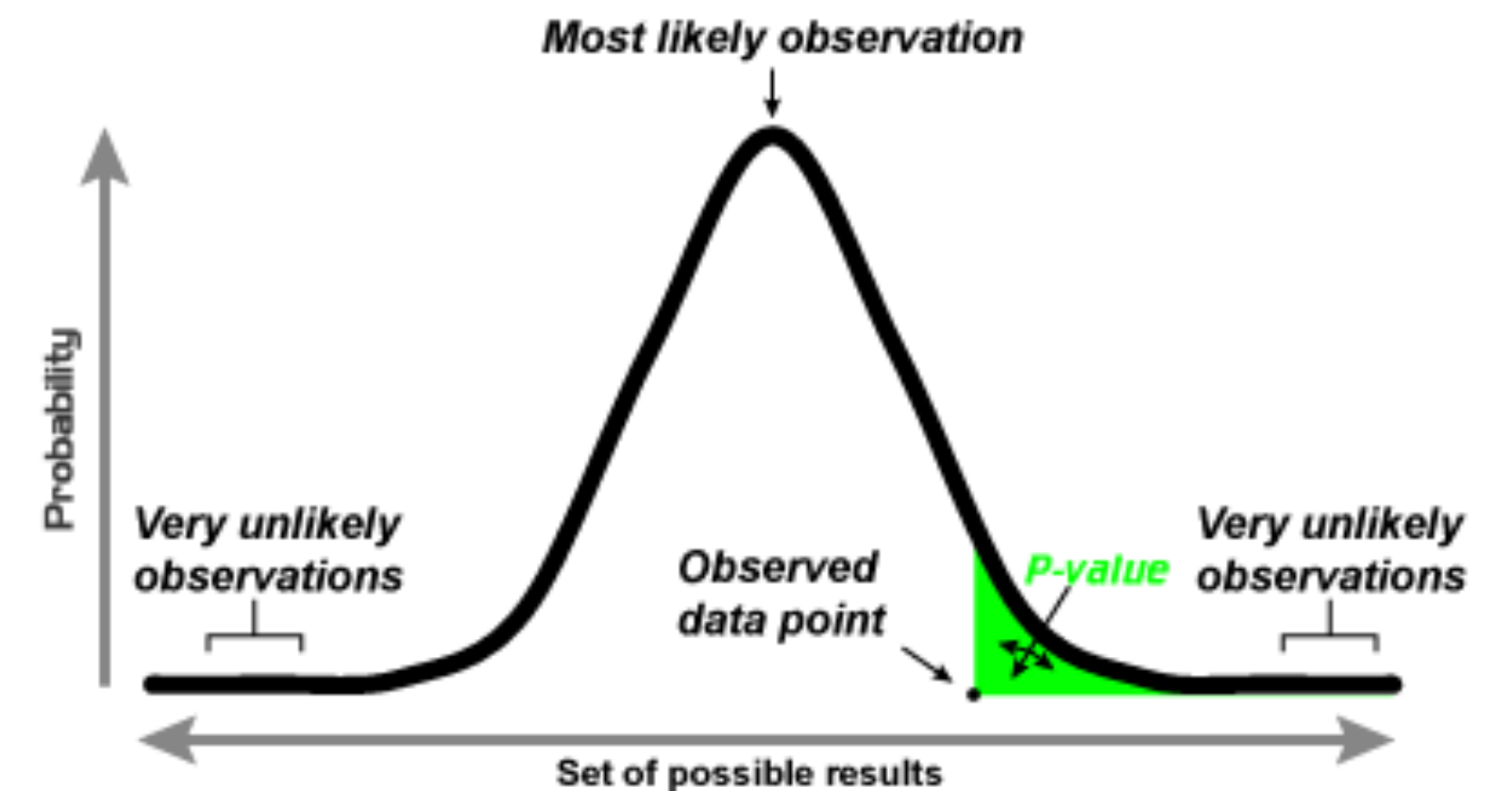
- Sometimes we are only interested in values departing from the mean in one direction
 - This is a **one-sided** or **one-tailed test**
- For example, suppose we want to assess whether our widgets are being produced at a significantly *higher* weight:

- $H_0 : \mu \leq 100g$

- $H_1 : \mu > 100g$

Null hypothesis is always the logical "opposite"

- How does the p -value compare between one and two-sided tests?



- Any given datapoint has *half* the p-value in a one-sided test than it does in a two-sided test
- We also do not divide α by 2 for a one-sided test, because all the area is now in one tail

simple extensions

- What do we do in a two-sample test when one of the samples violates the normal approximation assumptions?
 - Use a **two-sample t-test**
- Can we build a confidence interval around a mean when the normal approximation is violated?
 - Yes, as discussed, just use the t -statistic in place of the z -score
- What if we are only interested in a confidence interval on one side (e.g., a lower bound or an upper bound)?
 - Can use a **one-sided interval**, where one of the bounds is replaced by $-\infty$ or $+\infty$
 - When computing z_c or t_c , instead of $1 - (1 - c)/2$ (where dividing by 2), use $1 - (1 - c) = c$ since there is only one tail

